

# Setting up models

July 27, 2009

# Making models!

- Warm up with some things we've already covered here
- Look at more complicated models
- Take a close look at orthogonalization

# Model 1

- You measure BOLD activation for some task and you would like the group mean for this activation. You have 10 subjects, what would the design look like?

### General Linear Model

**EVs** | **Contrasts & F-tests**

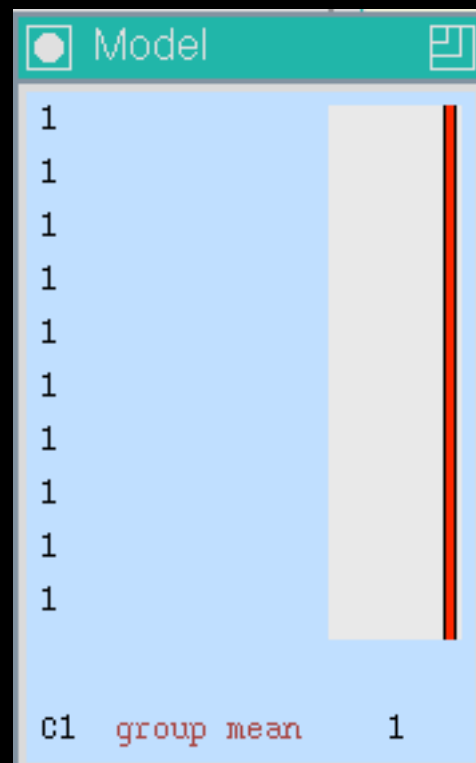
Number of main EVs: 1

Number of additional, voxel-dependent EVs: 0

Paste

	Group	EV1
Input 1	1	1
Input 2	1	1.0
Input 3	1	1.0
Input 4	1	1.0
Input 5	1	1.0
Input 6	1	1.0
Input 7	1	1.0
Input 8	1	1.0
Input 9	1	1.0
Input 10	1	1.0

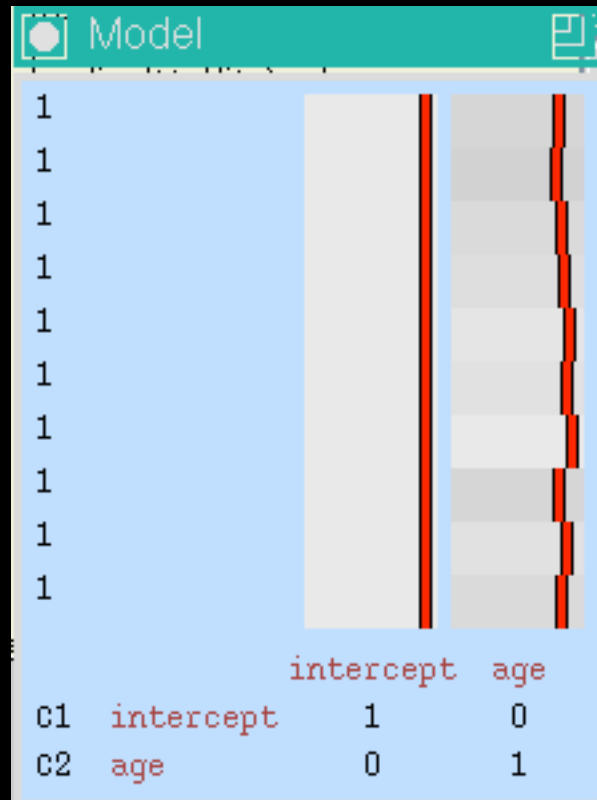
View design | Efficiency | Done



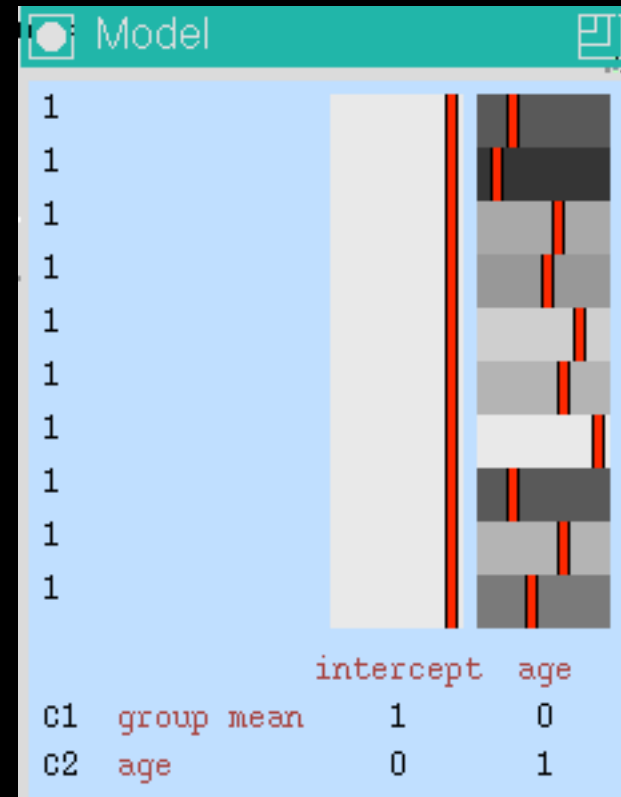
# Model 2

- Building off Model 1, you also have measured age and would like to see if there is an age effect.
  - What would the model look like?
  - What contrasts would you specify for the age effect?
  - Can you still obtain the overall mean from this model?

## Use age

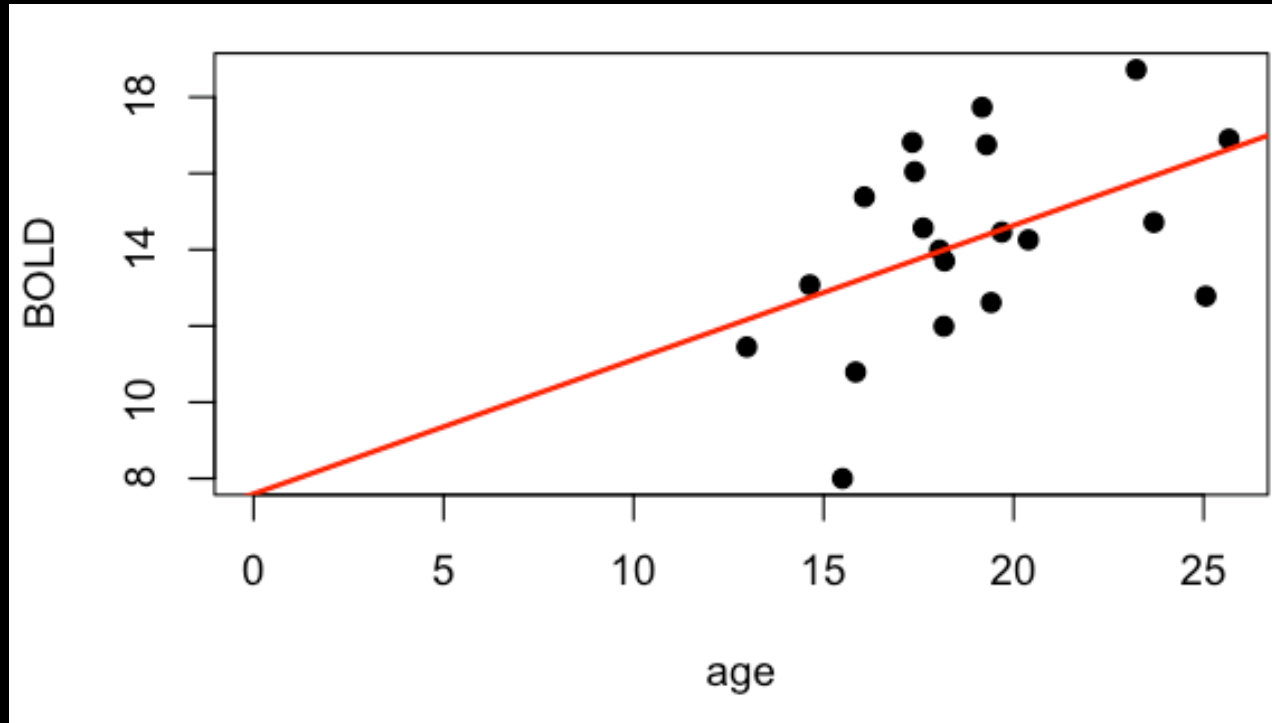


## Use demeaned age



- Both models will give exactly the same result for C2, but C1 will be different.

# Simulated data



Intercept ~ 7 or 8

Mean of BOLD ~ 14

# Simulated data

```
Call:
lm(formula = BOLD ~ age)

Residuals:
    Min       1Q   Median       3Q      Max
-5.04950 -1.36690 -0.01510  2.17838  3.38803

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  7.5955     3.0545   2.487  0.0229 *
age          0.3520     0.1594   2.207  0.0405 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1
```

Residual standard error: 2.359 on 18 degrees of freedom  
Multiple R-squared: 0.213, Adjusted R-squared: 0.1693  
F-statistic: 4.873 on 1 and 18 DF, p-value: 0.0405

```
Call:
lm(formula = BOLD ~ age_demeaned)

Residuals:
    Min       1Q   Median       3Q      Max
-5.04950 -1.36690 -0.01510  2.17838  3.38803

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  14.2367     0.5276  26.986 5.18e-16 ***
age_demeaned  0.3520     0.1594   2.207  0.0405 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1
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Models fit equally well

# Simulated data

```
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Coefficient for age and p-value don't change

# Simulated data

```
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Residuals:
    Min       1Q   Median       3Q      Max
-5.04950 -1.36690 -0.01510  2.17838  3.38803

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```

- Parameter estimate and p-value changes for “intercept”.
- In demeaned case, it is equal to the mean of BOLD

```
> mean(BOLD)
[1] 14.23671
```

# Summary of demeaning

- Only really necessary if you want your PE of column of 1s to be the overall mean
- Often people have rounding errors after demeaning. Double-check this when you do it.

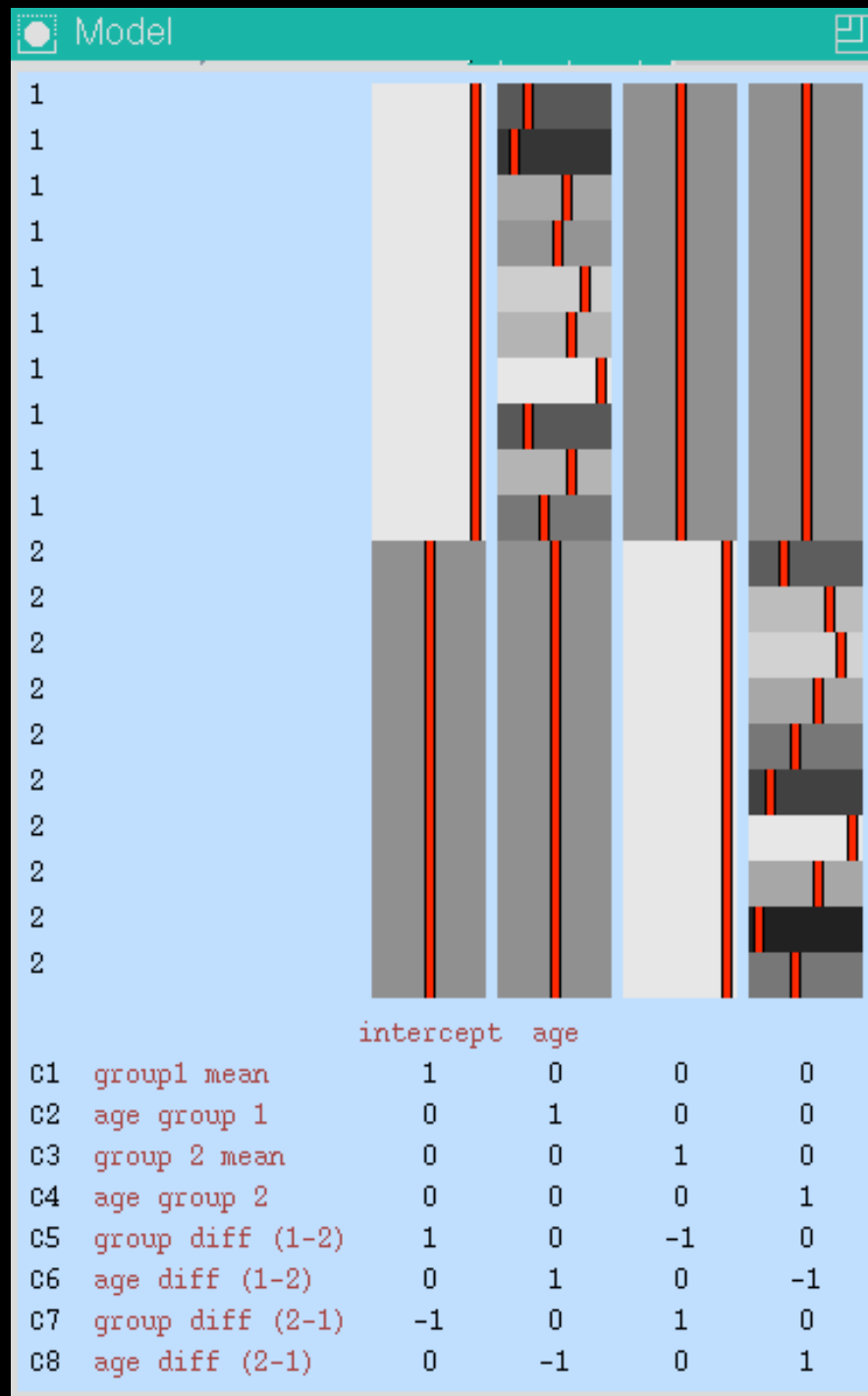
# Continuous covariates between groups

- You have 2 groups and a continuous measure for all subjects (age)
  - 2 things you can look at
    - Within group trends in age: What is the relationship of age within group?
    - Effect of age between groups: How do differences in age between groups impact group differences?

# Model 3

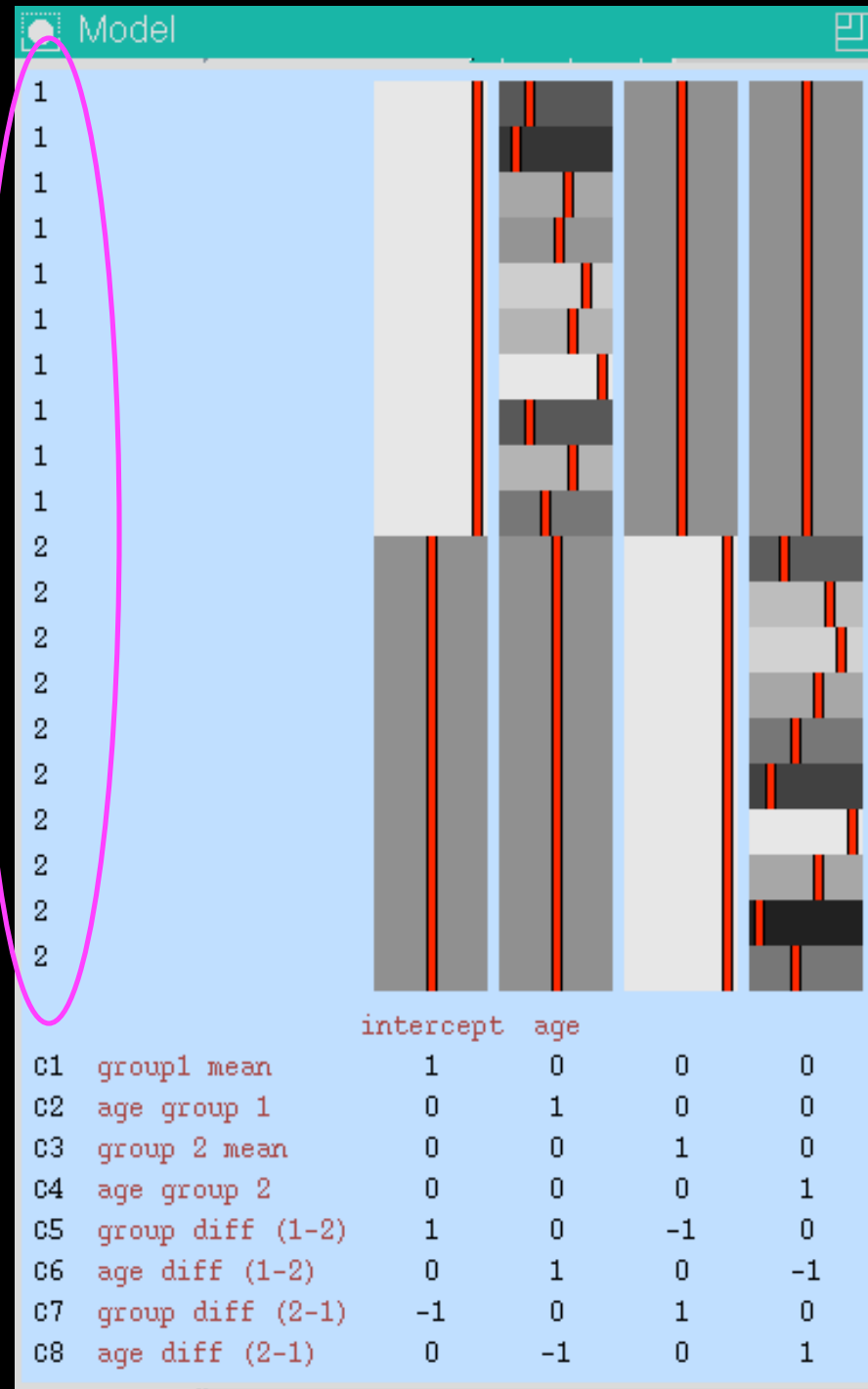
- Now you have 2 groups (10 sub each) and you'd like to see if there's a difference in overall mean between the groups and the within group age effect and how it compares between groups
  - What would the model look like?
  - What if your groups are expected to have different levels of variability (patients and controls)?
  - Should you demean age? How would you demean it?
  - What if the ages for your two groups are significantly different?

I demeaned age within group here since I'm interested in group means and not intercepts



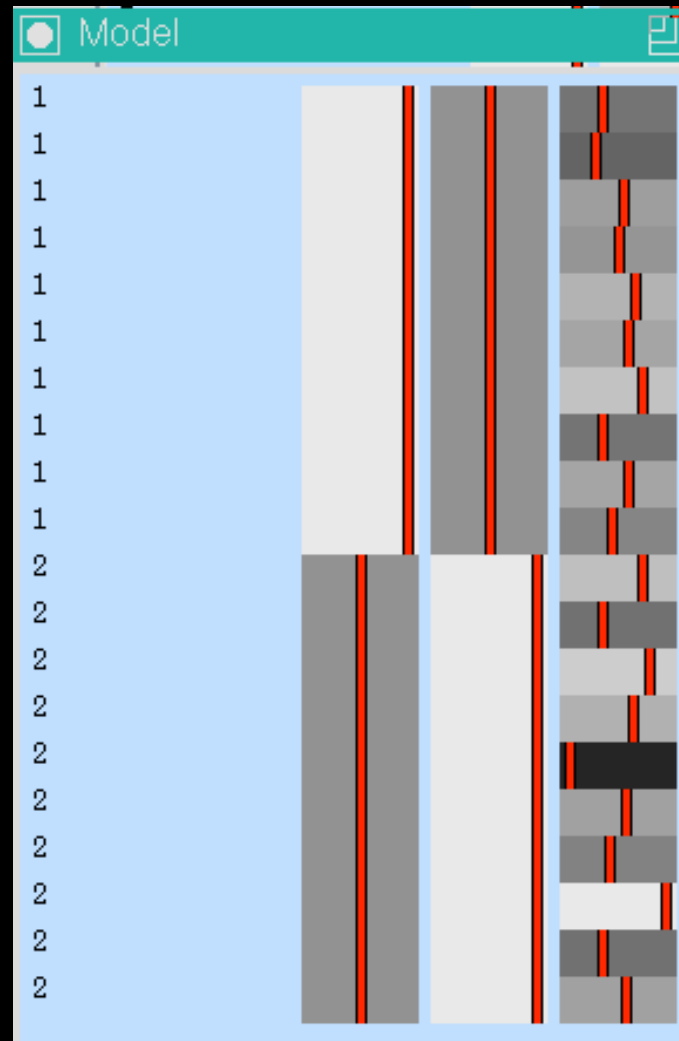
“group” column allows different variances in groups

Model must be separable!





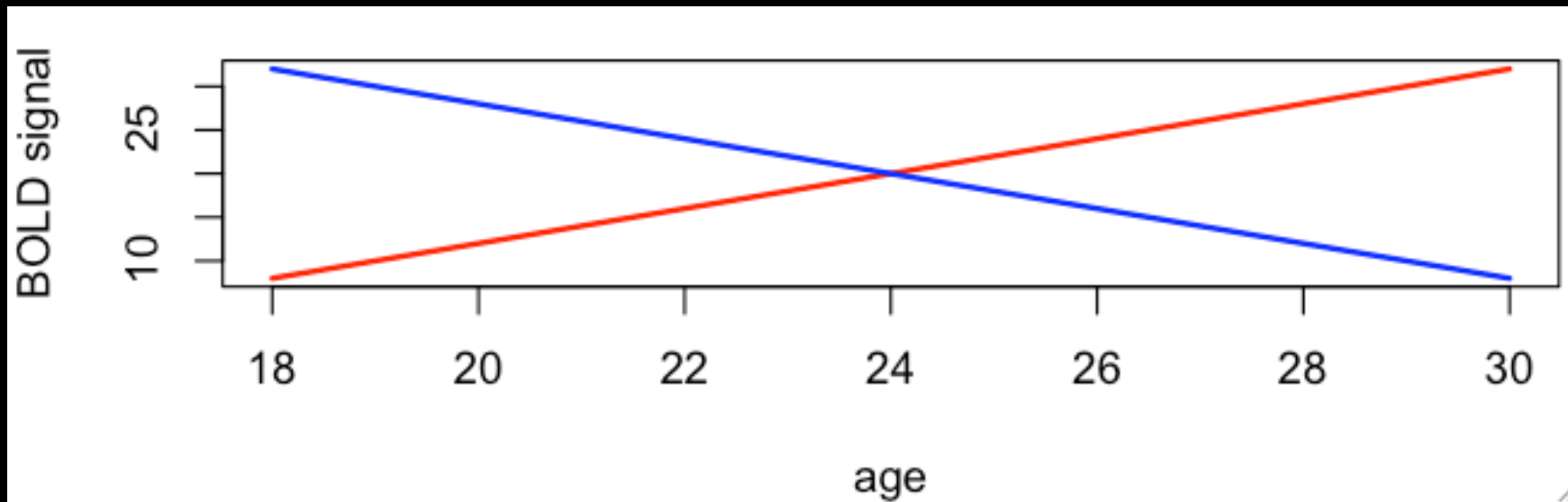
# Not separable



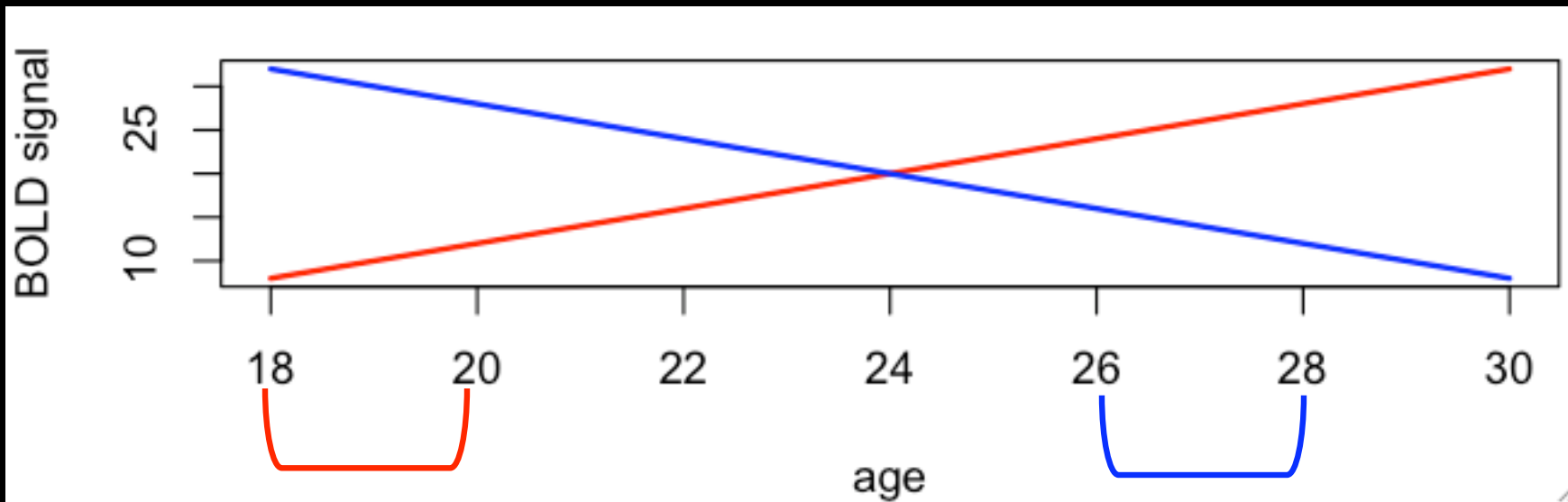
# What if age is different between groups?

- Differences in slopes may only be because you're sampling different parts of the distribution
- Must assume trend is linear

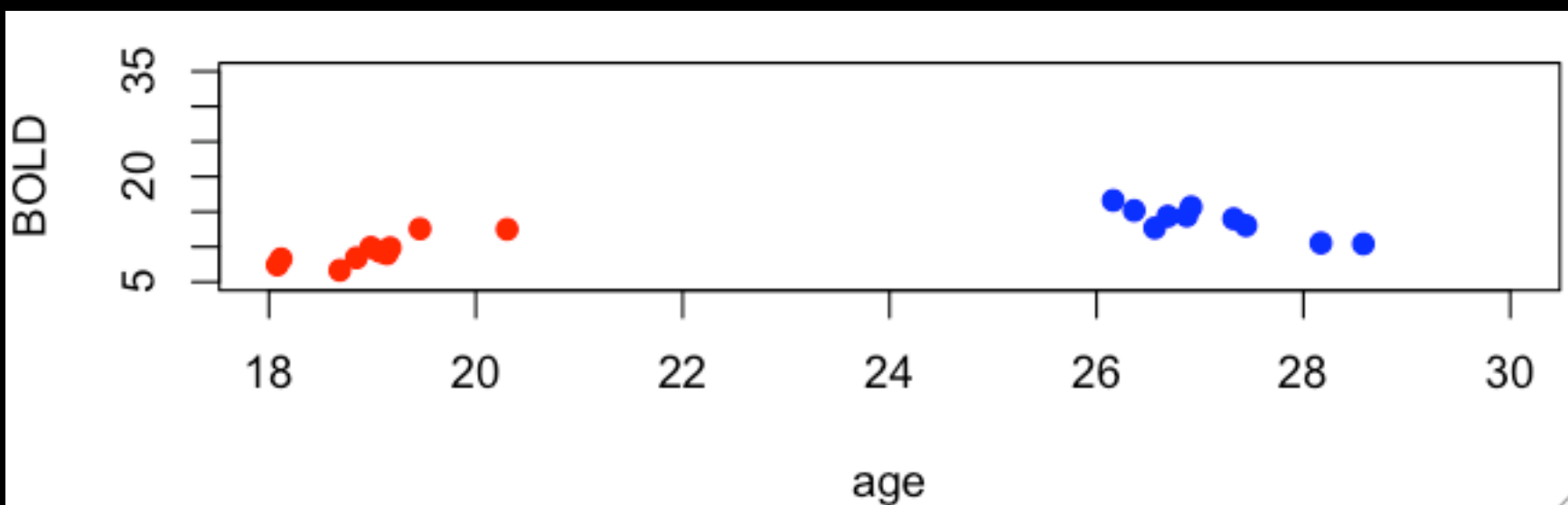
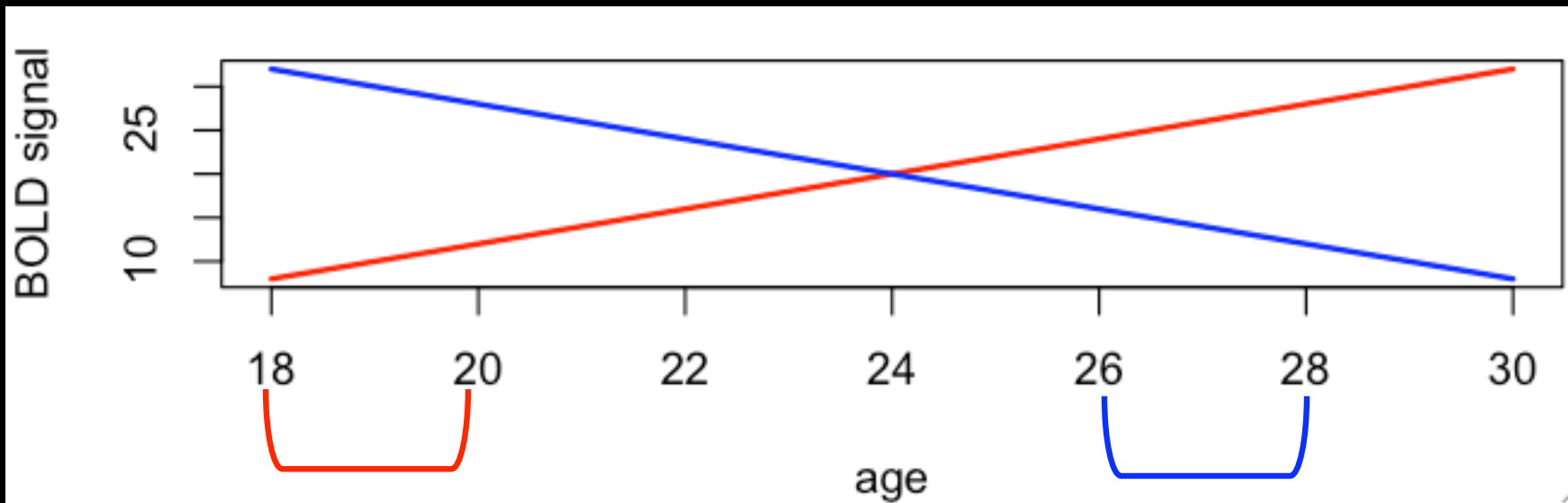
When trends are linear, it doesn't matter where I sample age, I get the same slope



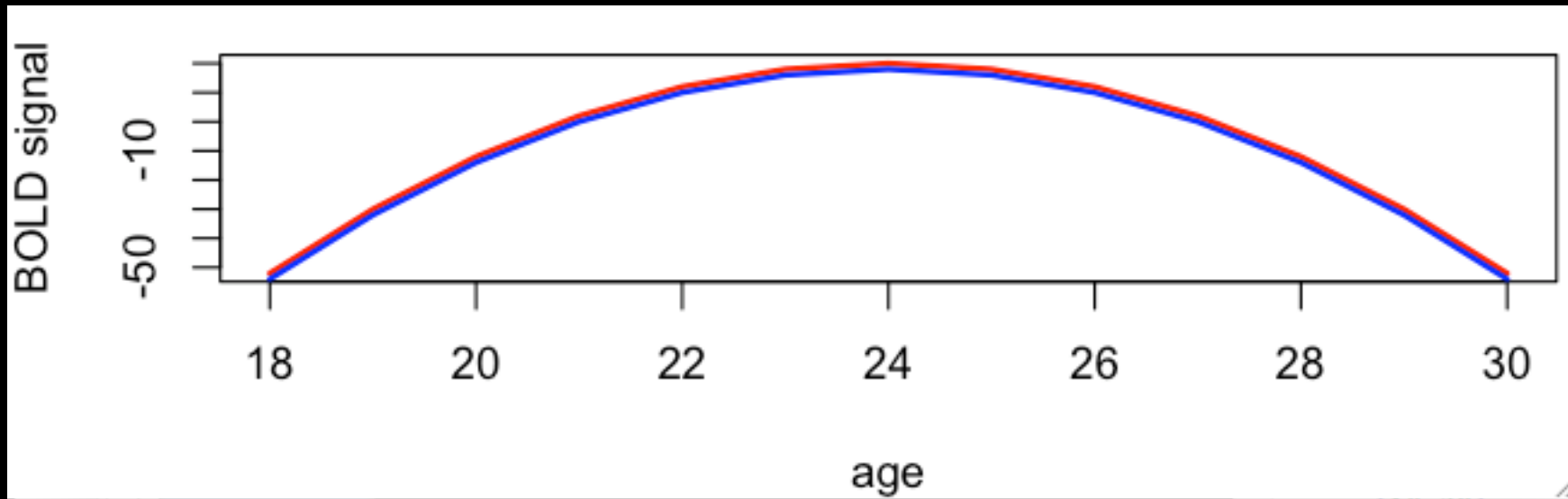
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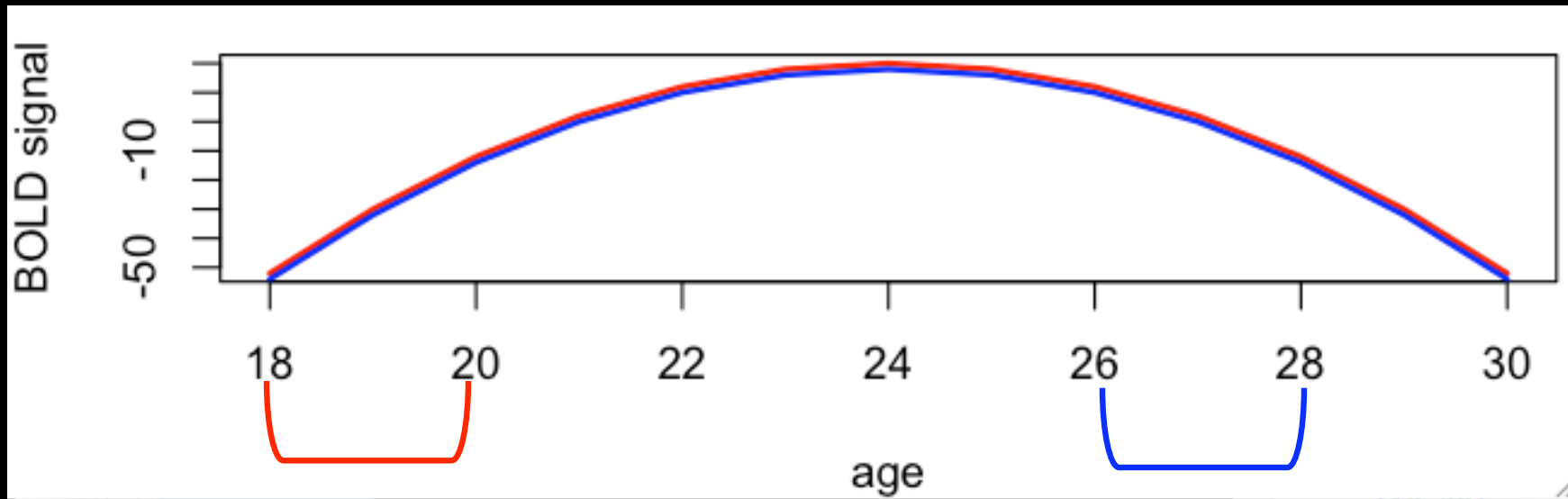
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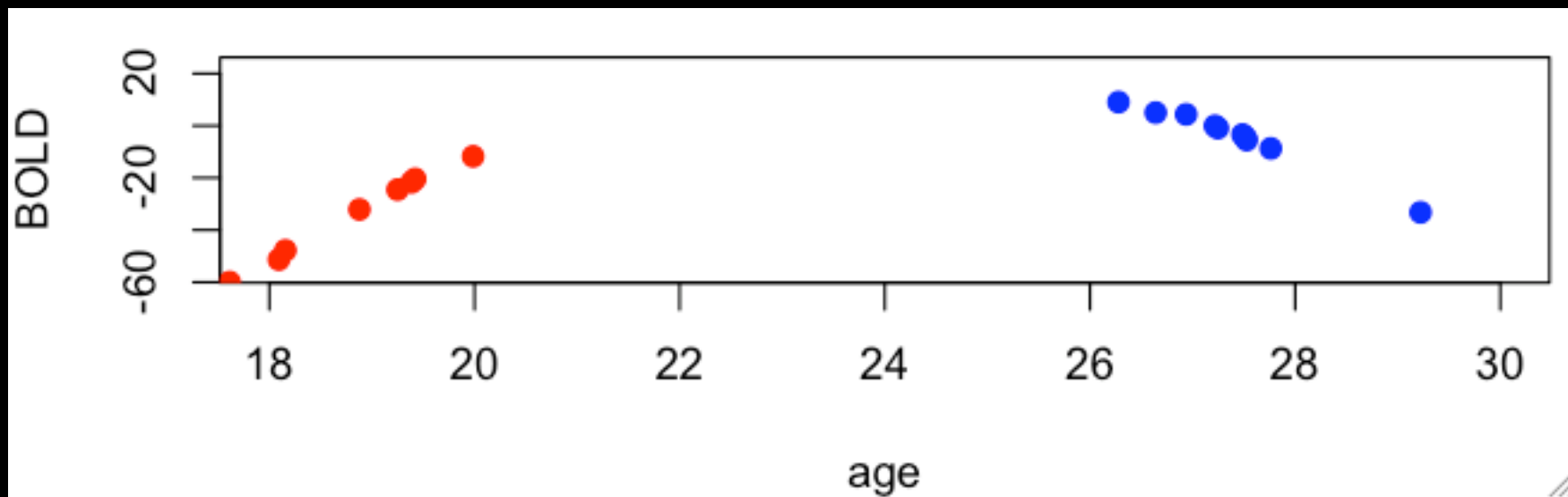
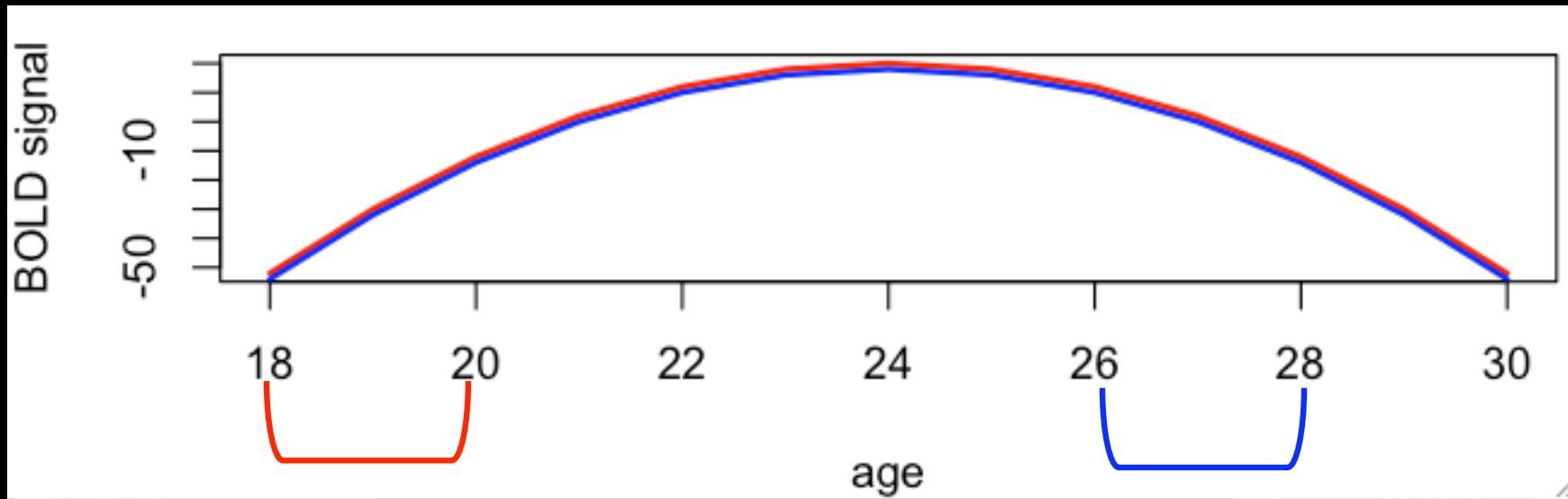
With a nonlinear trend, the trend may be the same in both groups, but sampling different ranges show different trends



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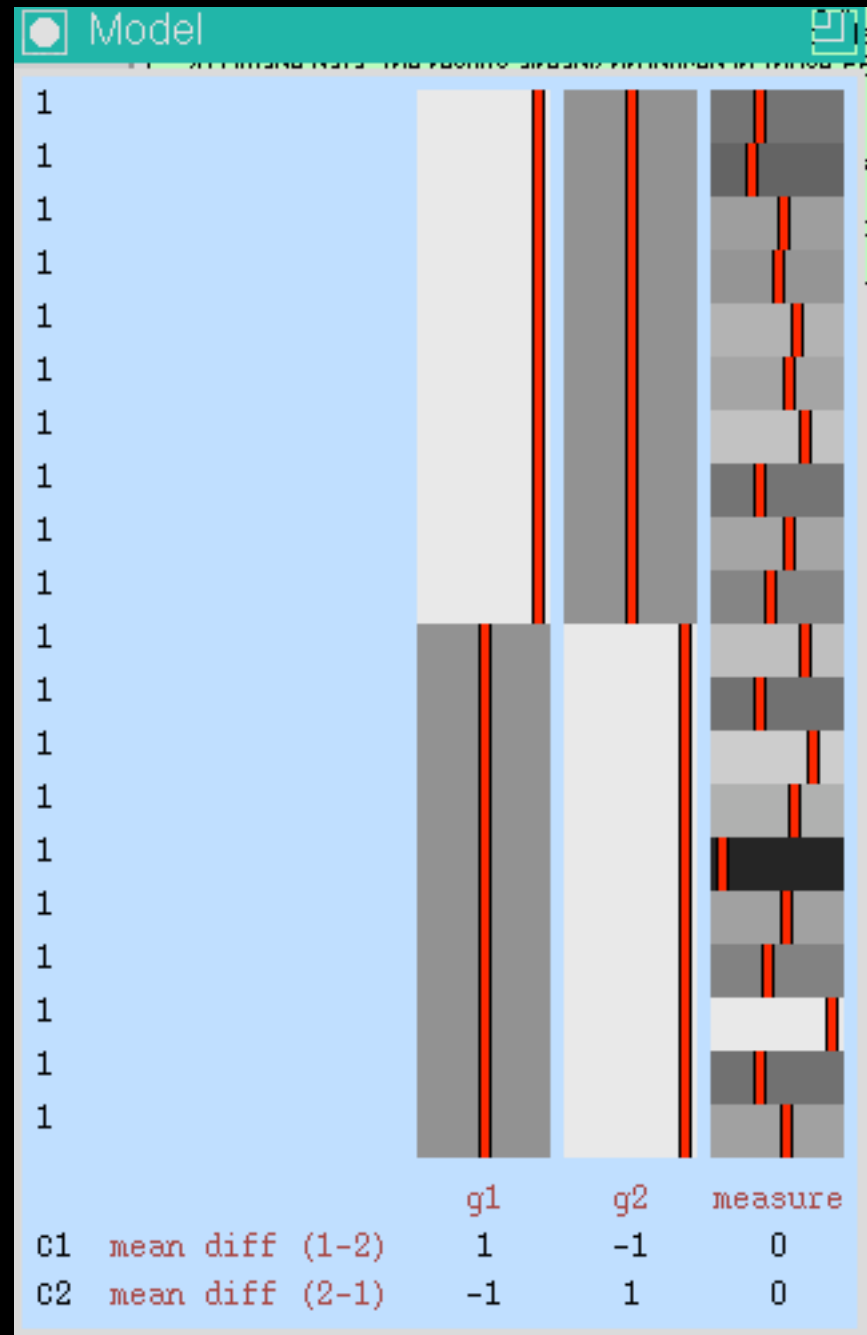




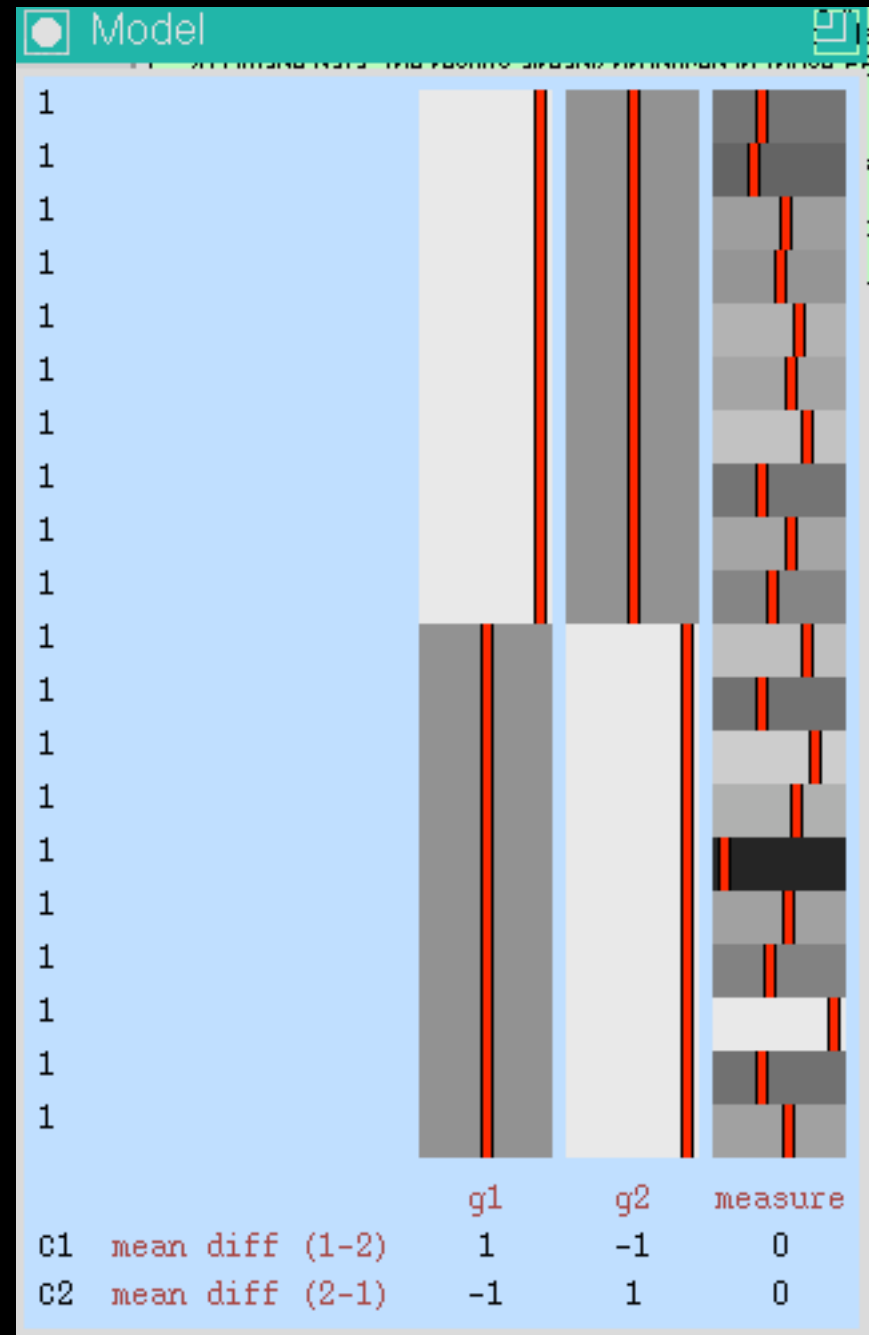
# Model 4

- Similar to Model 3 we have two groups and a confounding covariate (maybe BIS or age). Our primary interest is in the difference of means between the two groups.
  - What is the model to simply test the difference in means
  - If I wanted to make sure this difference wasn't due to between group differences in BIS, what would that model look like?
  - What are some restrictions of this model?

- Do not demean the confounding measure \*especially\* within group
  - This removes any confounding effect the measure might have

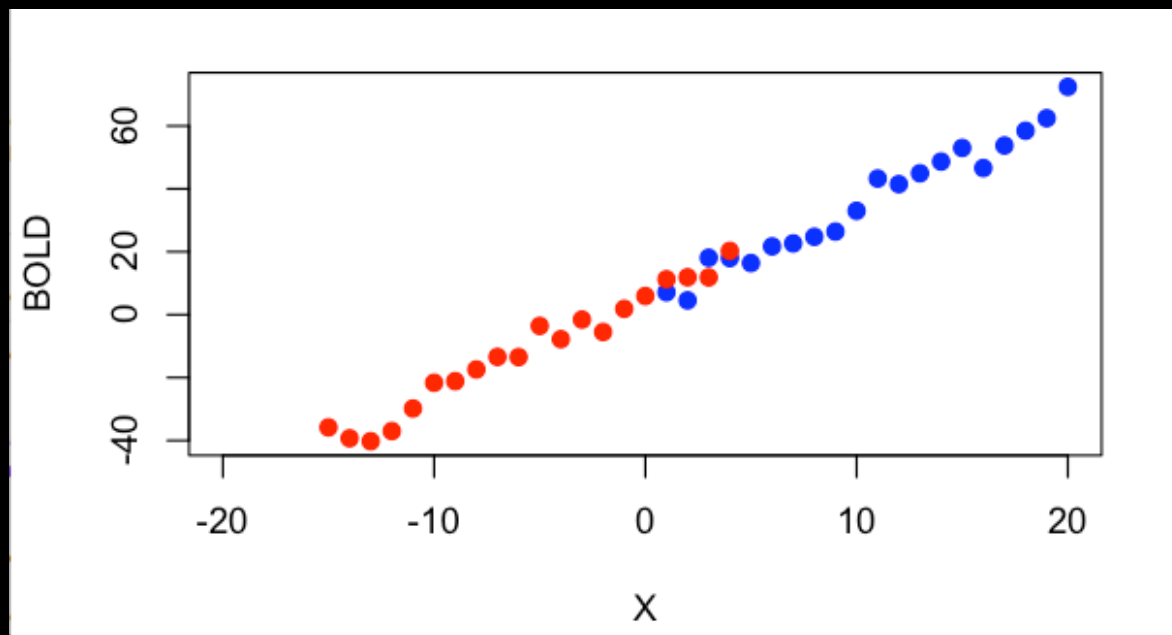


- Do not demean the confounding measure \*especially\* within group
  - This removes any confounding effect the measure might have
- Restriction: Cannot create a separable model, so cannot have different group variances.
  - Separable model only adjusts for the measure within group and we're interested in between group differences here.



# Why you shouldn't demean in this case

- What if this is what your data look like?
  - Difference in means is clearly due to range of  $X$  sampled, not the group membership



# Model comparisons

- Group difference
  - Model had column of 1s and group1 indicator

```
Call:
lm(formula = y ~ group)

Residuals:
    Min       1Q   Median       3Q      Max
-31.3533 -15.0814  0.6285  14.1005  36.5253

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  -11.232     4.255  -2.640  0.0120 *
group         47.124     6.018   7.831 1.90e-09 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1

Residual standard error: 19.03 on 38 degrees of freedom
Multiple R-squared: 0.6174, Adjusted R-squared: 0.6073
F-statistic: 61.32 on 1 and 38 DF, p-value: 1.903e-09
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```

Groups are significantly different  
( $p < 0.0001$ )

# Model comparisons

- Add in confounder (X) correctly

```
lm(formula = y ~ group + x)

Residuals:
    Min       1Q   Median       3Q      Max
-6.6737 -2.1005 -0.2804  1.8046  6.4766

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  6.16477    0.94590   6.517 1.26e-07 ***
group       -3.48417    1.89667  -1.837  0.0743 .
x           3.16302    0.09617  32.892 < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1

Residual standard error: 3.507 on 37 degrees of freedom
Multiple R-squared:  0.9873, Adjusted R-squared:  0.9867
F-statistic: 1444 on 2 and 37 DF, p-value: < 2.2e-16
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```

Overall fit  
has  
improved  
( $R^2=.61$  in  
last model)



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```

Group diff =  
-3.48417

Not  
significant

# Model comparisons

- Separate age and demean within group

No demeaning

```
lm(formula = y ~ group + x_g1 + x_g2)

Residuals:
    Min       1Q   Median       3Q      Max
-6.8453 -1.9445 -0.2159  1.9138  6.1802

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)    5.9932     1.0971   5.463 3.63e-06 ***
group          -3.6402     1.9808  -1.838  0.0744 .
x_g1           3.1942     0.1377  23.201 < 2e-16 ***
x_g2           3.1318     0.1377  22.748 < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.55 on 36 degrees of freedom
Multiple R-squared:  0.9874, Adjusted R-squared:  0.9863
F-statistic: 939.2 on 3 and 36 DF,  p-value: < 2.2e-16
```

With demeaning

```
lm(formula = y ~ group + x_g1_dm + x_g2_dm)

Residuals:
    Min       1Q   Median       3Q      Max
-6.8453 -1.9445 -0.2159  1.9138  6.1802

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  -11.2319     0.7939  -14.15 2.82e-16 ***
group         47.1242     1.1227   41.97 < 2e-16 ***
x_g1_dm       3.1942     0.1377   23.20 < 2e-16 ***
x_g2_dm       3.1318     0.1377   22.75 < 2e-16 ***
---
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# Model comparisons

- Separate age and demean within group

No demeaning

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Coefficients:
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(Intercept)  5.9932     1.8971   3.163 0.0026 ***
group       -3.6402     1.9808  -1.838  0.0744 .
x_g1        3.1942     0.1377  23.201 < 2e-16 ***
x_g2        3.1318     0.1377  22.748 < 2e-16 ***
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lm(formula = y ~ group + x_g1_dm + x_g2_dm)

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    Min       1Q   Median       3Q      Max
-6.8453 -1.9445 -0.2159  1.9138  6.1802

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 11.2310     0.7020  15.992 < 2e-16 ***
group       47.1242     1.1227  41.97 < 2e-16 ***
x_g1_dm     3.1942     0.1377  23.20 < 2e-16 ***
x_g2_dm     3.1318     0.1377  22.75 < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

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```

Group difference is quite large after demeaning

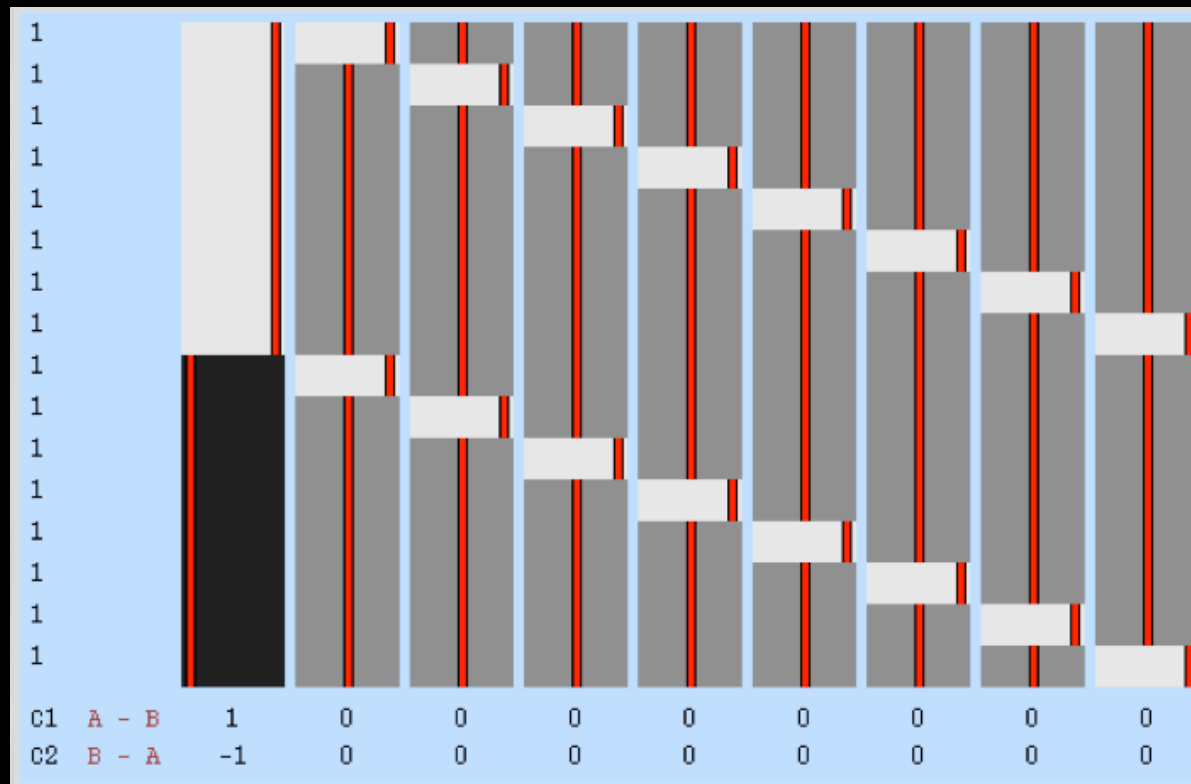
# Differences between 3 & 4

- In model 3 we were interested in testing the differences in trends between two groups.
  - age modeled to show within group trends, which are then compared across groups
- In model 4 we were only interested in comparing the means between two groups and whether it was due to a confounder.
  - X modeled to show between group differences

# Model 4

- You have estimated successful stop pre and post training (2 runs) for each subject. At the group level you have 8 subjects with 2 measures each, what is the appropriate model?

# Paired t test



# Model 5

- You have 9 subjects and all but one subject have complete pairs of data (one subject missed the second scan)
  - Can you still use a paired t test?
  - Other options?

# Model 5

- Suggestions
  - If there is high within subject correlation (like in the tire example the other day) you need a paired test
    - Toss the subject with incomplete data
  - If there isn't a high correlation, a two-sample t-test may be okay
    - Be careful or else you'll be losing power
    - Use complete data to compare paired t test to 2 sample t test



# Model 6

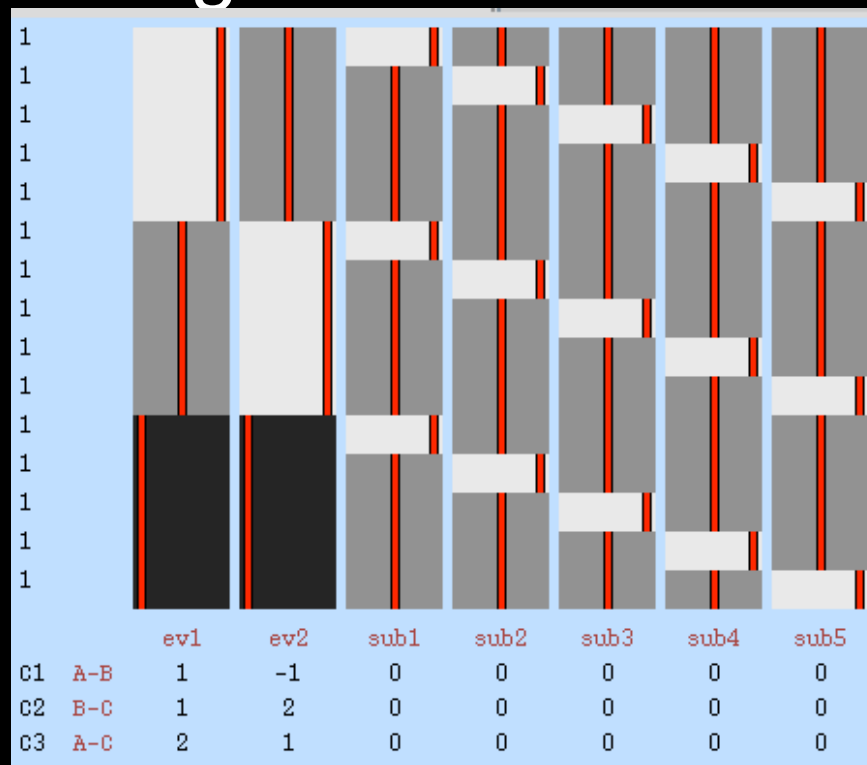
- You have 5 subjects scanned under 3 conditions (A, B & C) and you want to make all pairwise comparisons in one model
  - You must account for repeated measures
  - Construct 3 contrasts: A-B, B-C, A-C

# Triple paired t test

	ev1	ev2	sub1	sub2	sub3	sub4	sub5
A1	1	0	1	0	0	0	0
A2	1	0	0	1.0	0	0	0
A3	1	0	0	0	1.0	0	0
A4	1	0	0	0	0	1.0	0
A5	1	0	0	0	0	0	1.0
B1	0.0	1.0	1	0	0	0	0
B2	0.0	1.0	0	1.0	0	0	0
B3	0.0	1.0	0	0	1.0	0	0
B4	0.0	1.0	0	0	0	1.0	0
B5	0.0	1.0	0	0	0	0	1.0
C1	-1.0	-1.0	1	0	0	0	0
C2	-1.0	-1.0	0	1.0	0	0	0
C3	-1.0	-1.0	0	0	1.0	0	0
C4	-1.0	-1.0	0	0	0	1.0	0
C5	-1.0	-1.0	0	0	0	0	1.0

# Triple paired t test

- Really just a repeated measures ANOVA using factor effects



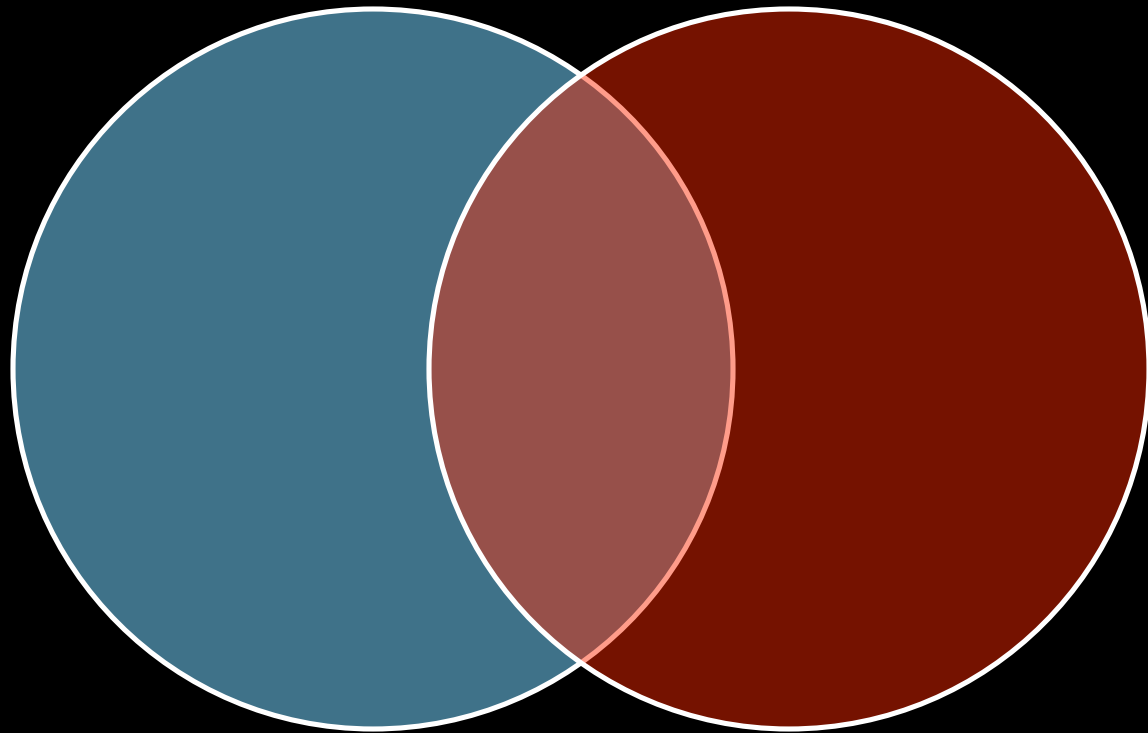
# Note on repeated measures ANOVA

- As long as each subject has repeated measures for all cells of the ANOVA use factor effects and split up the intercept into subject-specific means
- Won't work if not repeated across both factors
  - Eg. 2x2 ANOVA
    - Factor 1: Pre/post training (everybody has both)
    - Factor 2: control/unhealthy control (each subject only in one group or the other)

# Orthogonalization...what is it?

- Demeaning is the simplest form of orthogonalization
  - Orthogonalizes your regressor to the column of 1s (earlier example)
- Generally A orth wrt B
  - What is left over in A after regressing out B

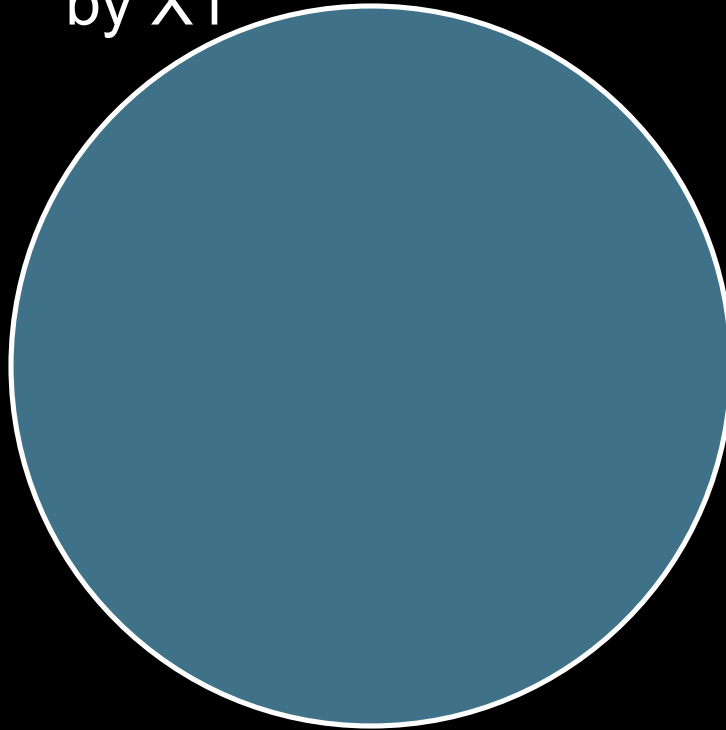
# How the GLM works



Variability in Y

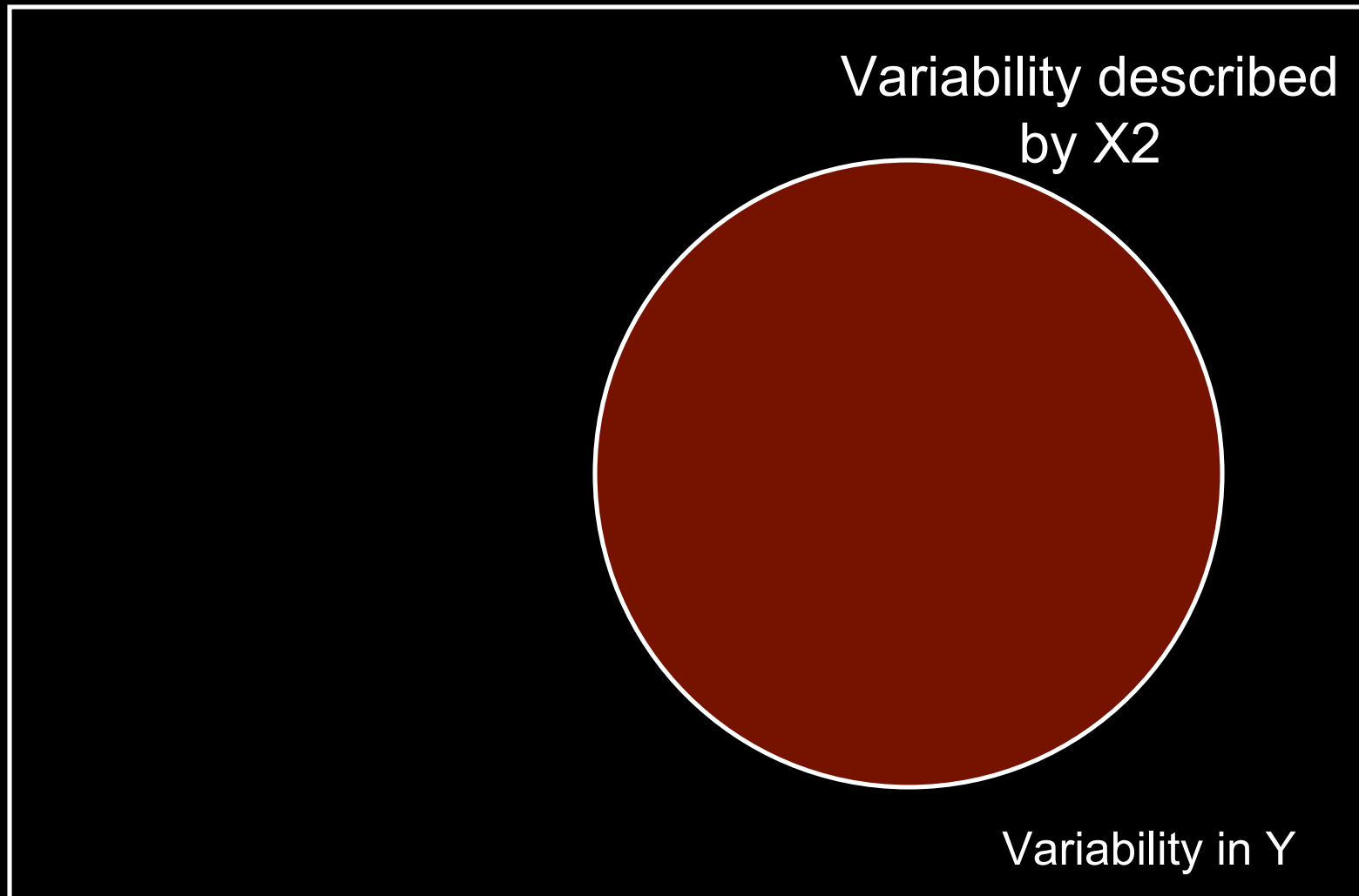
# How the GLM works

Variability described  
by  $X_1$



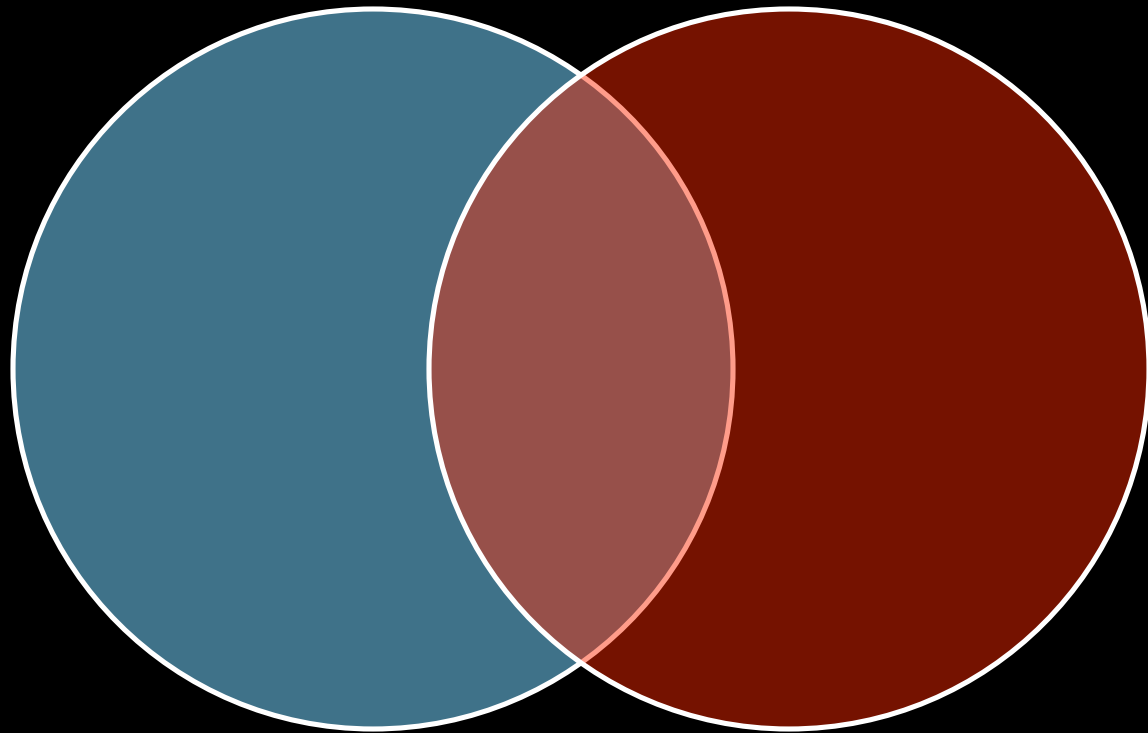
Variability in Y

# How the GLM works



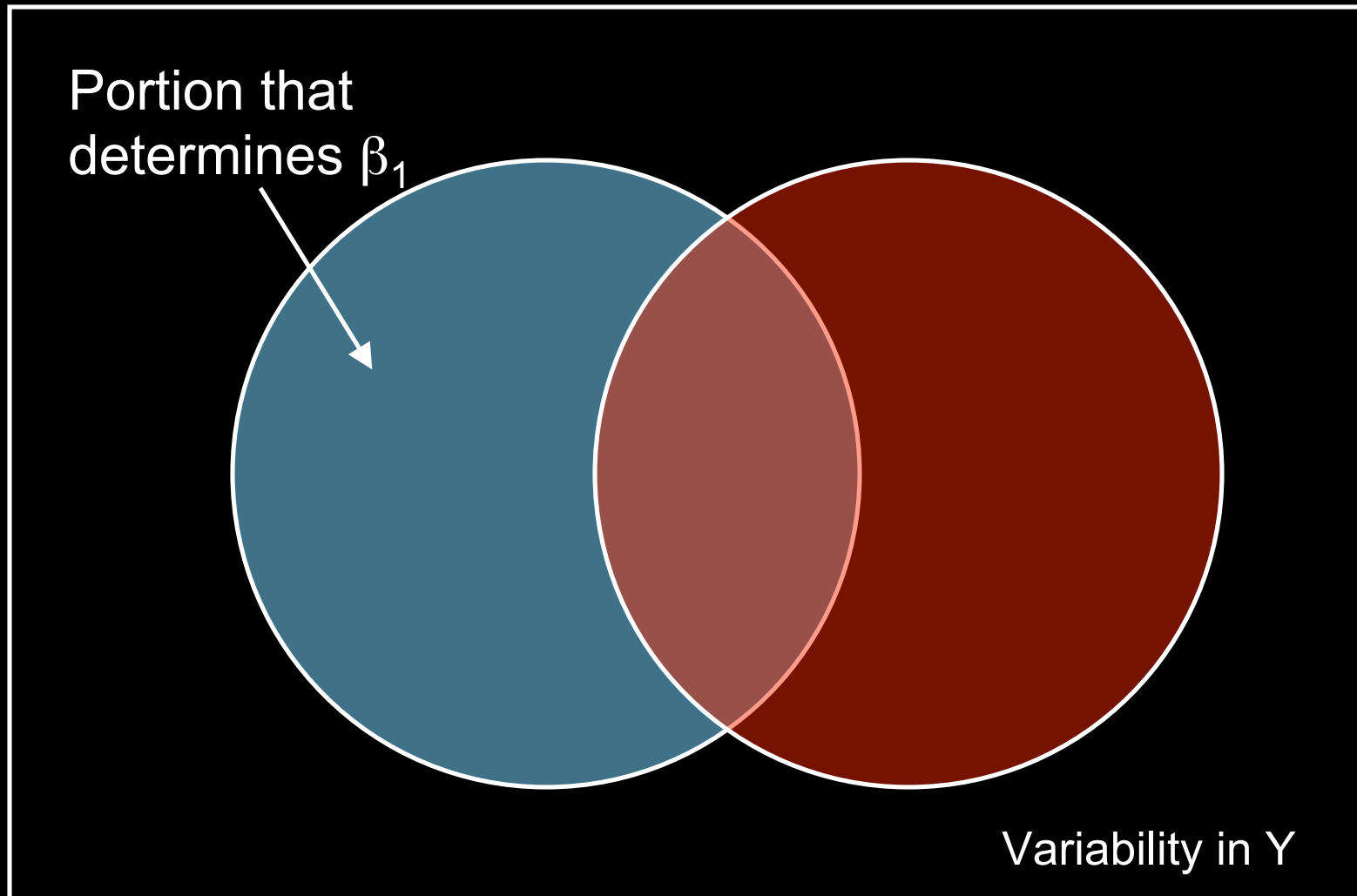


# How the GLM works

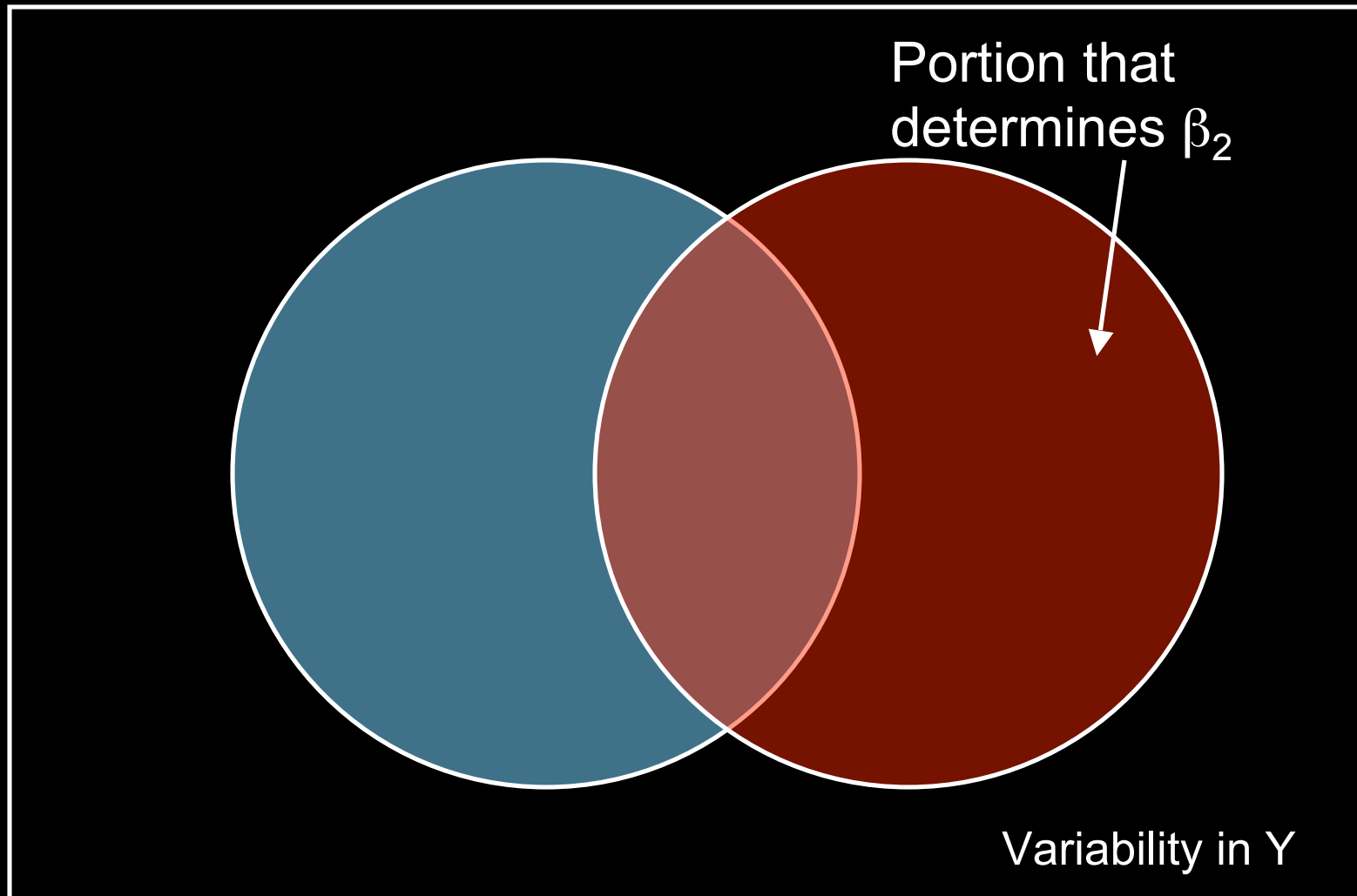


Variability in Y

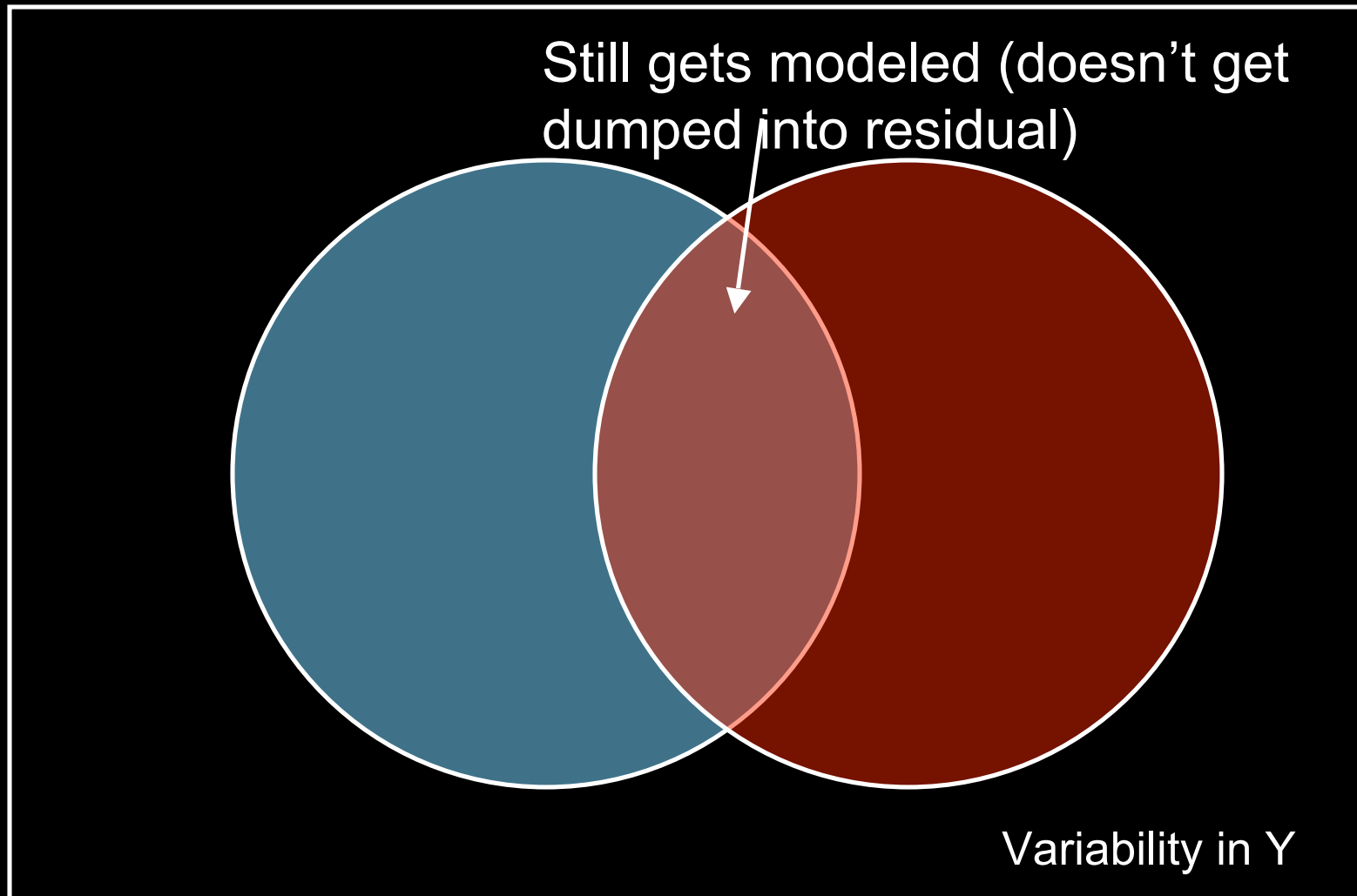
# How the GLM works



# How the GLM works

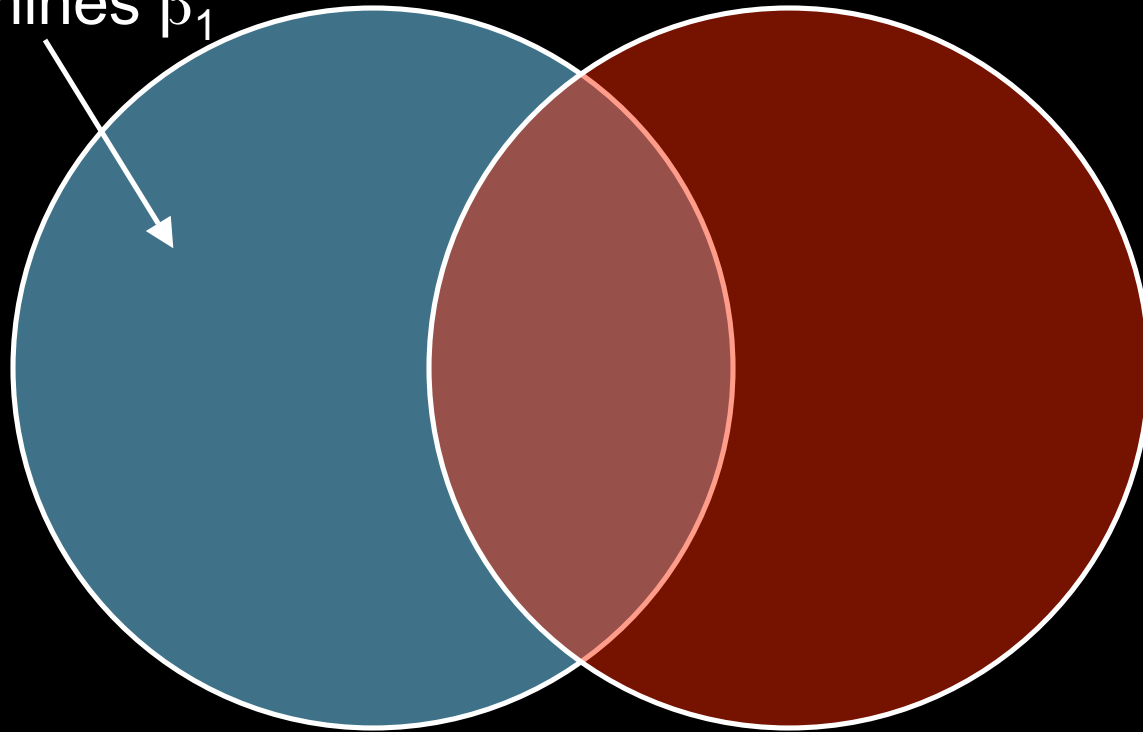


# How the GLM works



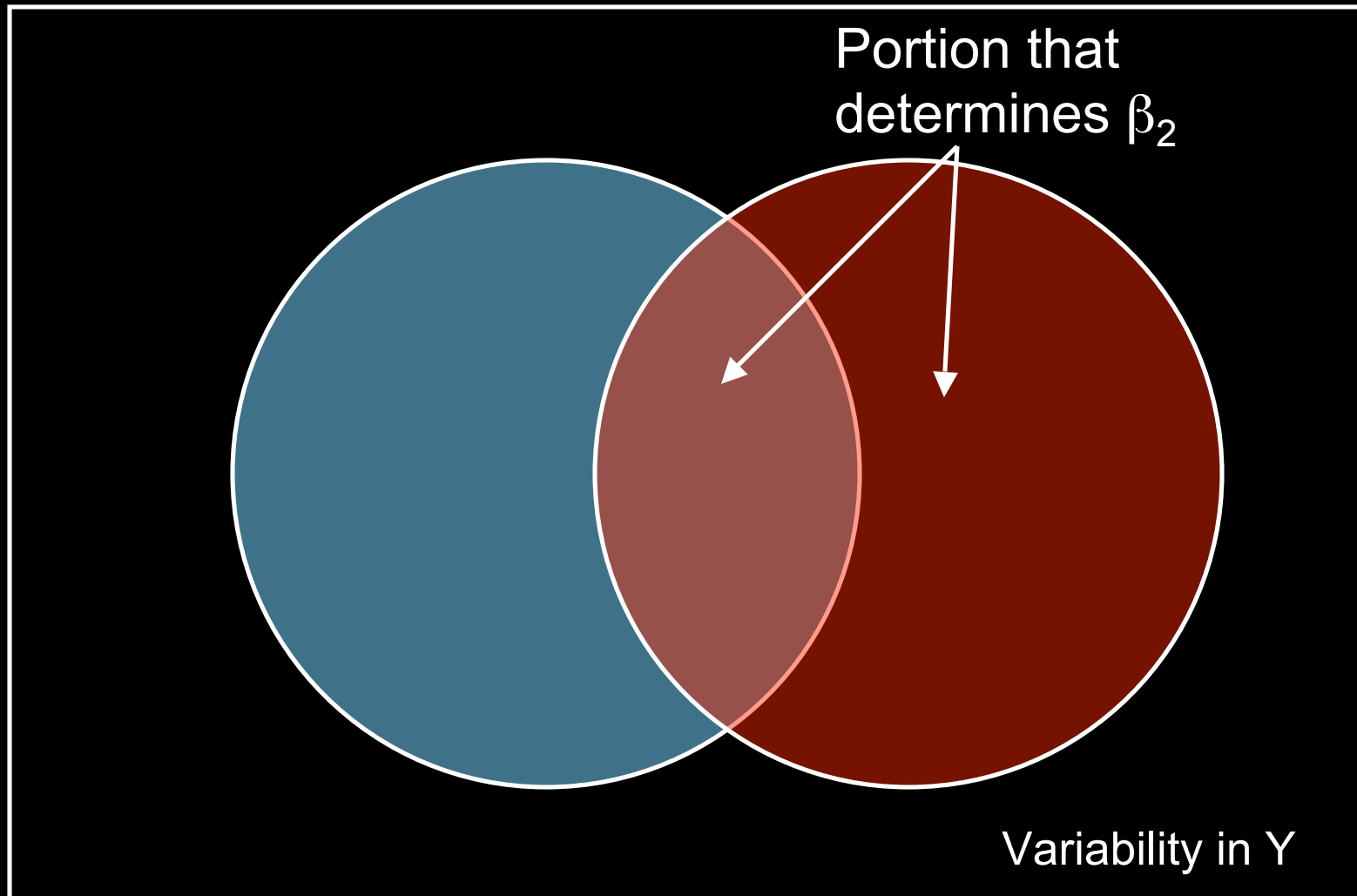
# Orthogonalize X1 wrt X2

Portion that  
determines  $\beta_1$



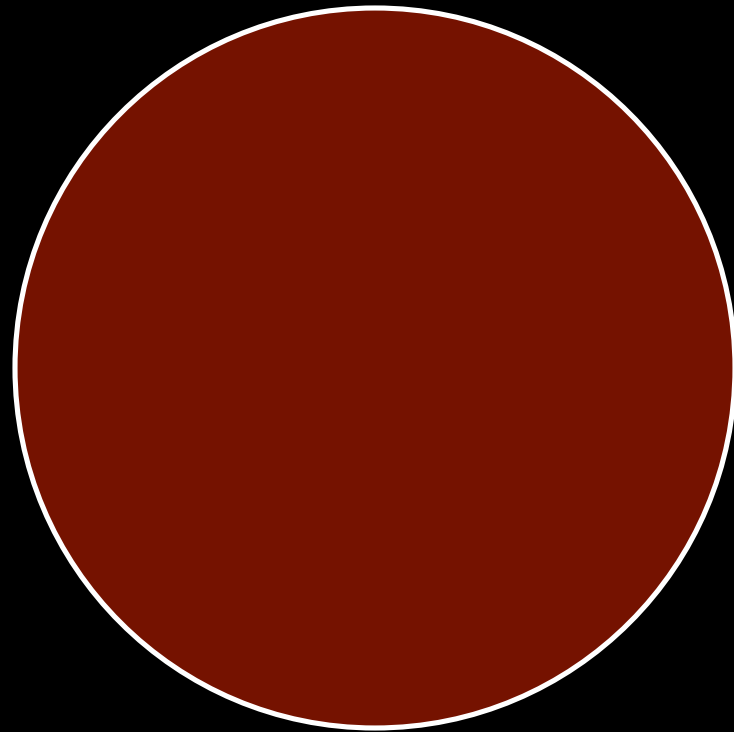
Variability in Y

# Orthogonalize $X_1$ wrt $X_2$



# Orthogonalize $X_1$ wrt $X_2$

It is like you are dealing with 2 independent things



Variability in  $Y$

# You can visualize regression using vectors

$$Y = X_1\beta_1 + X_2\beta_2$$

$$\begin{pmatrix} 3 \\ 4 \\ 0 \end{pmatrix} = \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} \beta_1 + \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix} \beta_2$$

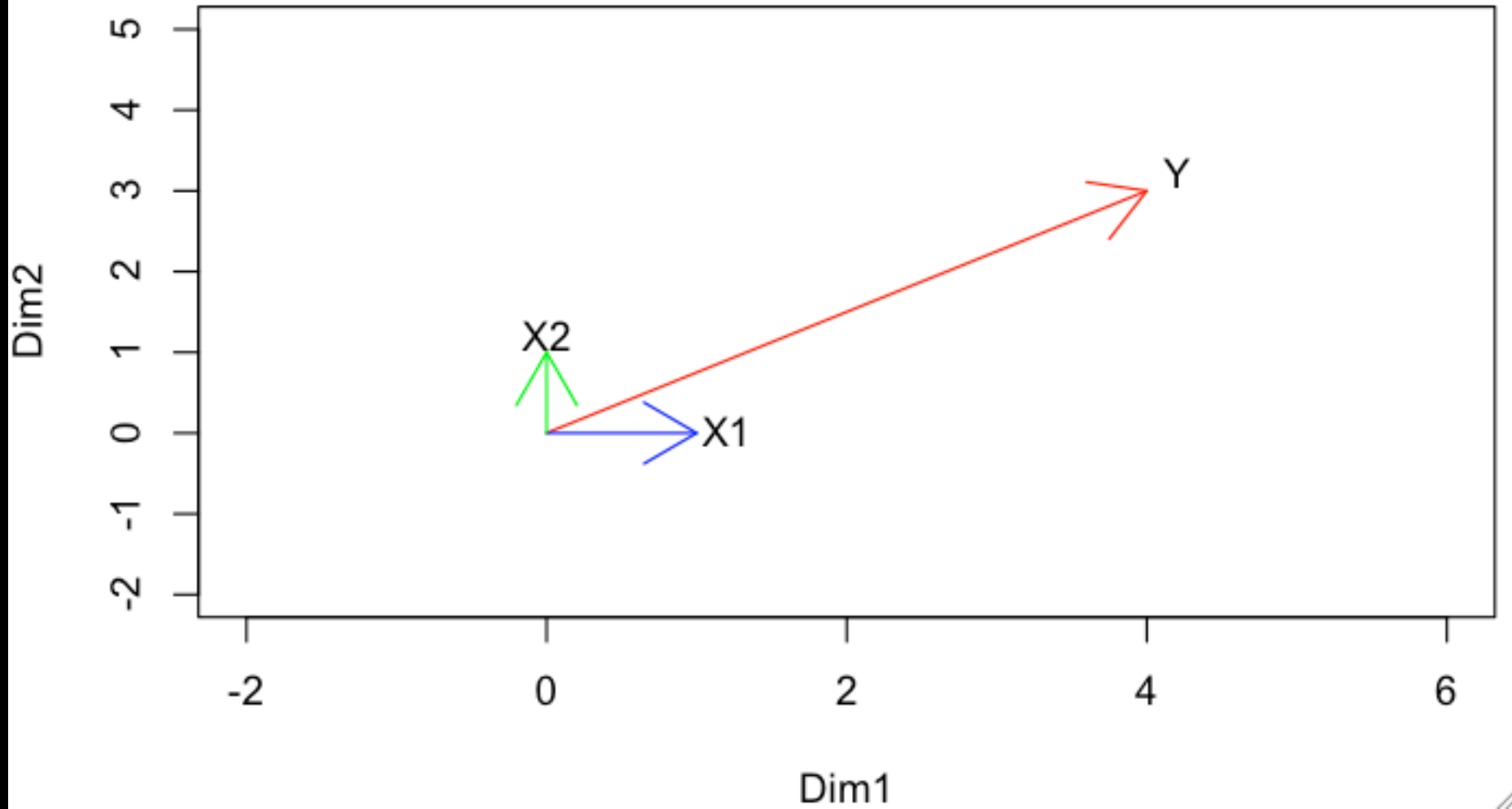


# You can visualize regression using vectors

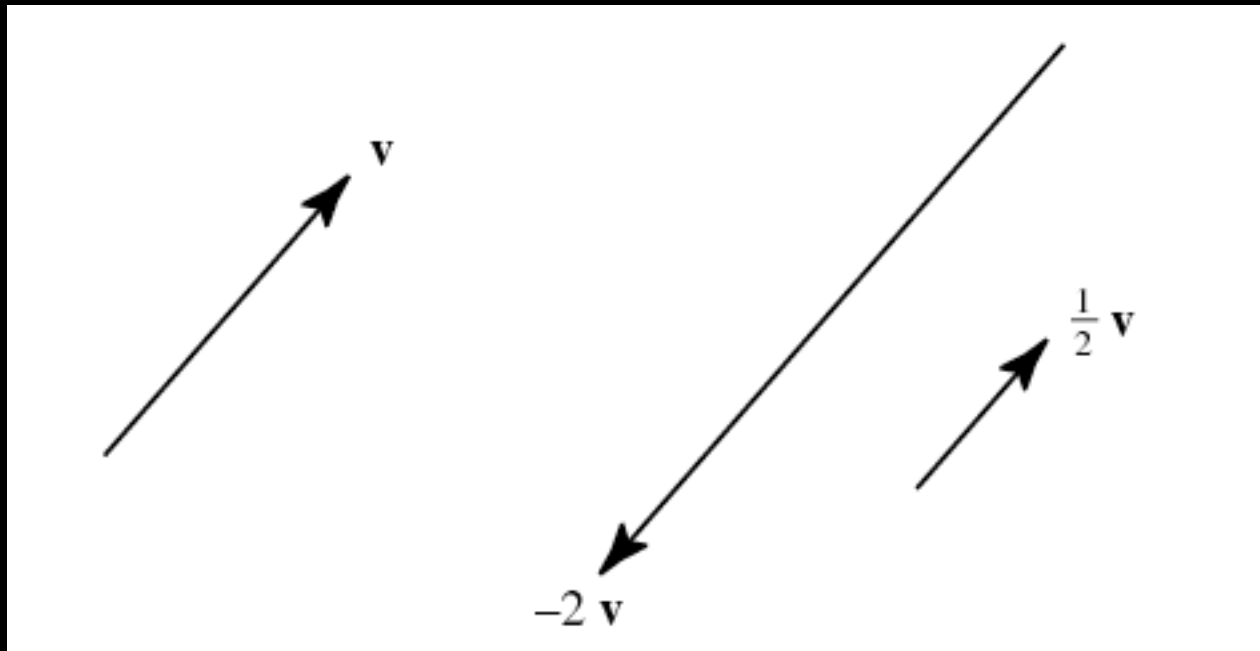
$$Y = X_1\beta_1 + X_2\beta_2$$

$$\begin{pmatrix} 3 \\ 4 \\ 0 \end{pmatrix} = \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} \beta_1 + \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix} \beta_2$$

$$\beta_1=3 \text{ and } \beta_2=4$$

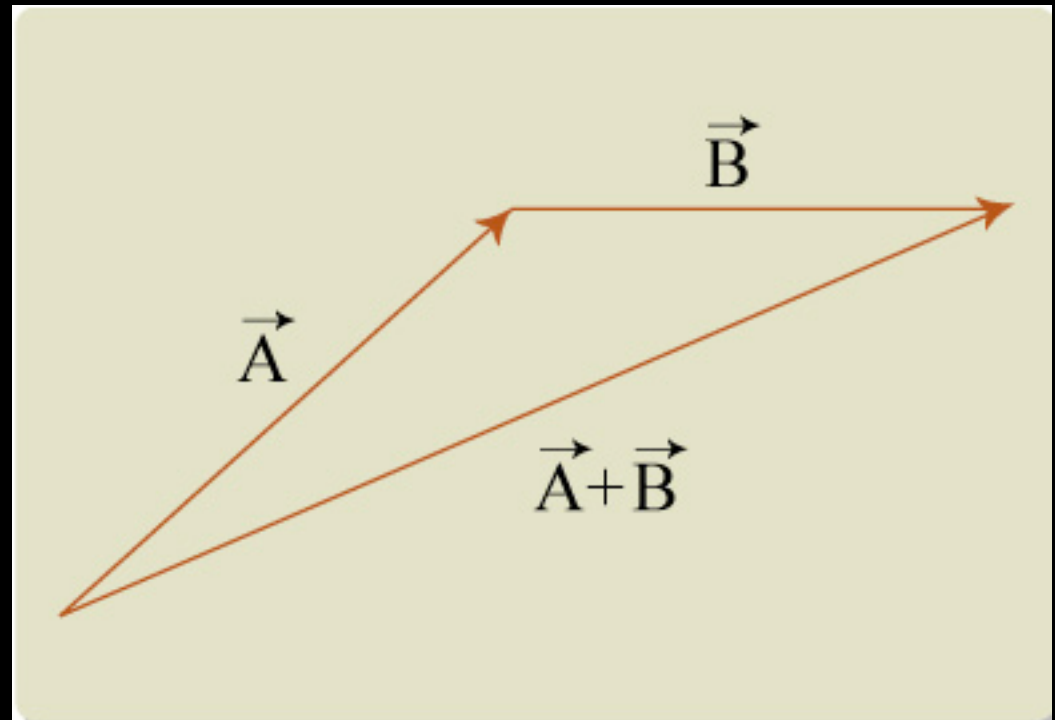


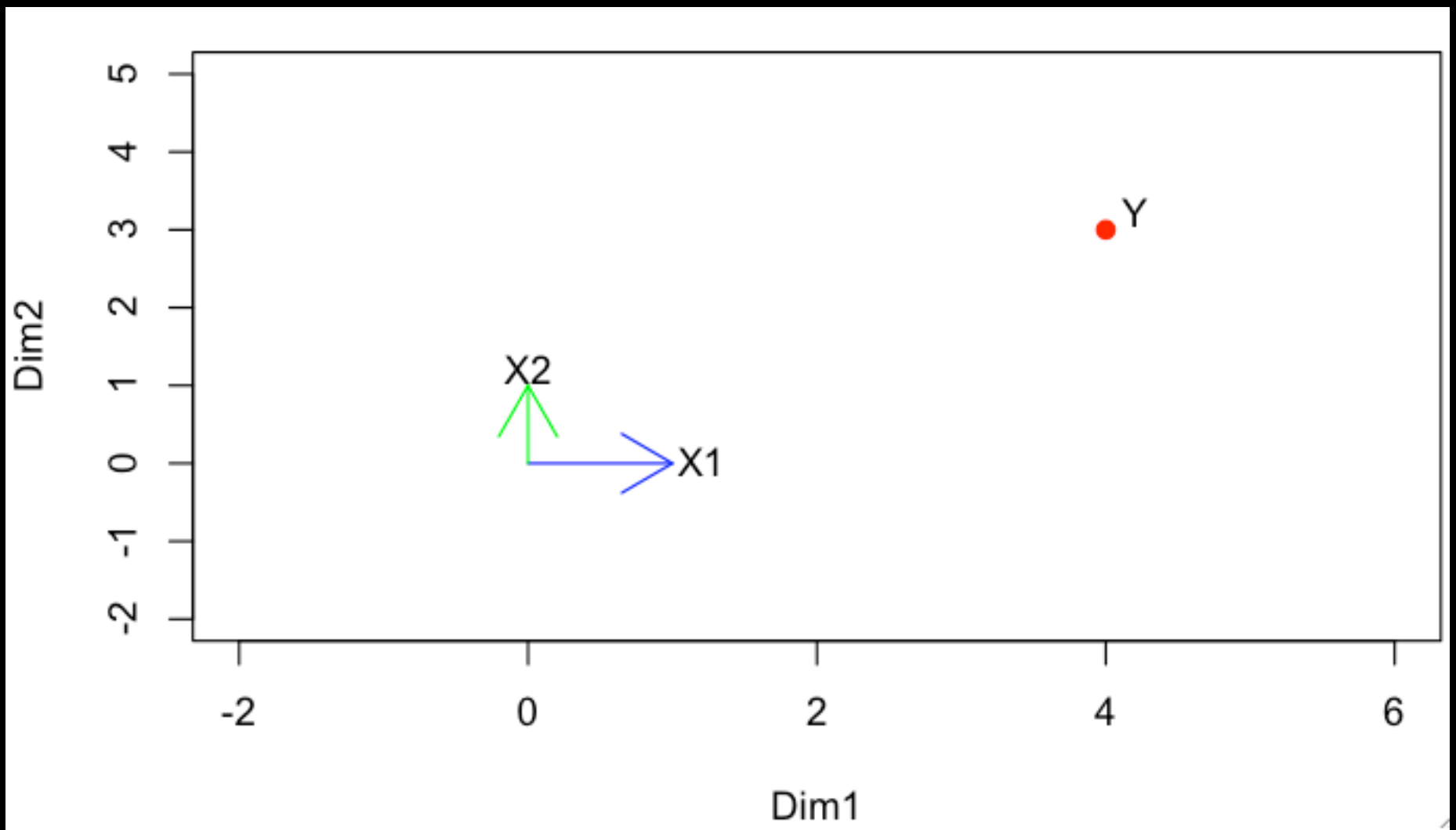
# Scalar multiplication of a vector

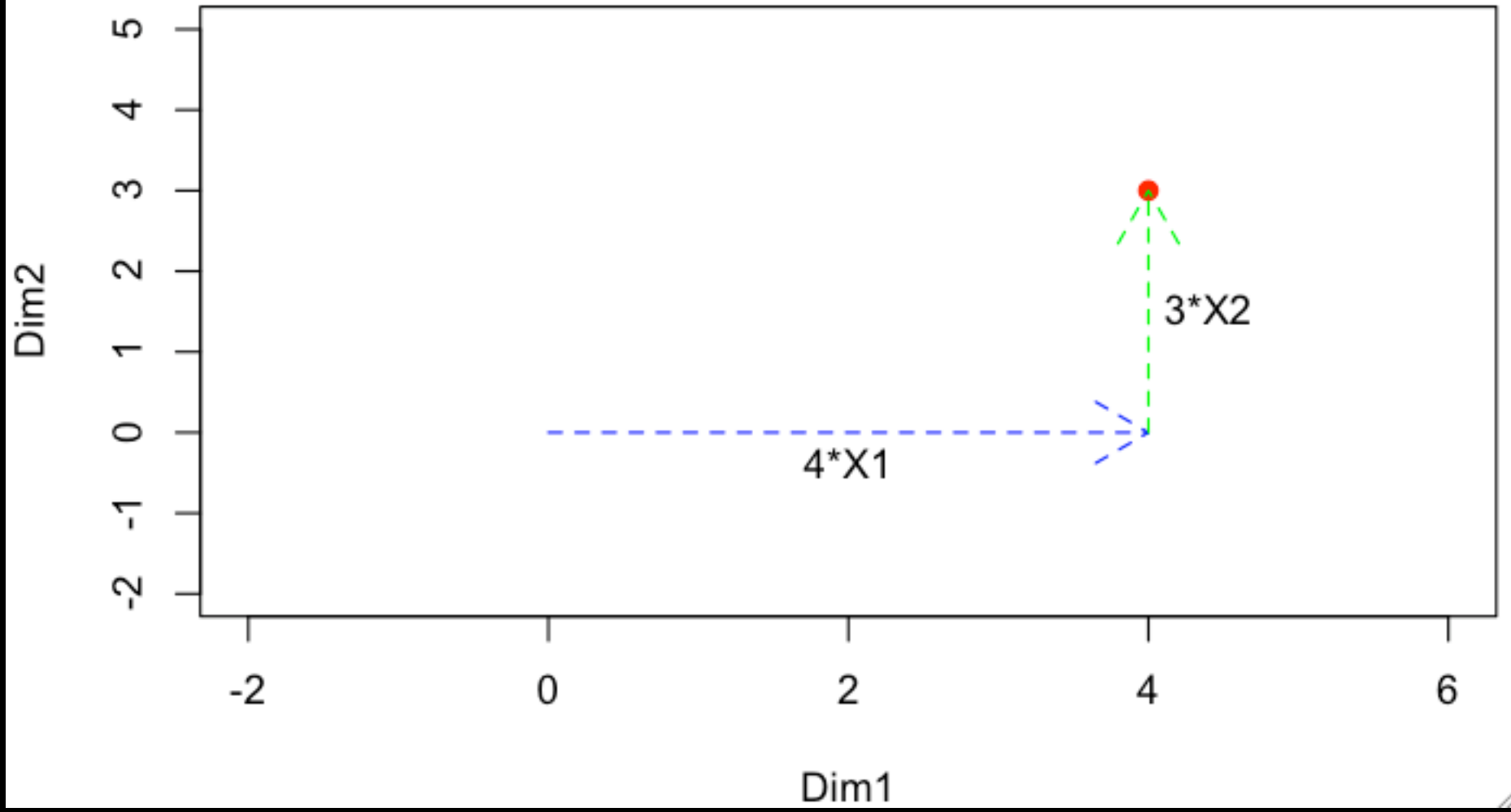


# Vector addition

To add, put tail  
of one vector  
at head of  
another vector







# Non-orthogonal case

$$Y = X_1\beta_1 + X_2\beta_2$$

$$\begin{bmatrix} 4 \\ 5 \\ 0 \end{bmatrix} = \begin{bmatrix} 2 \\ 2 \\ 0 \end{bmatrix} \beta_1 + \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix} \beta_2$$

# Non-orthogonal case

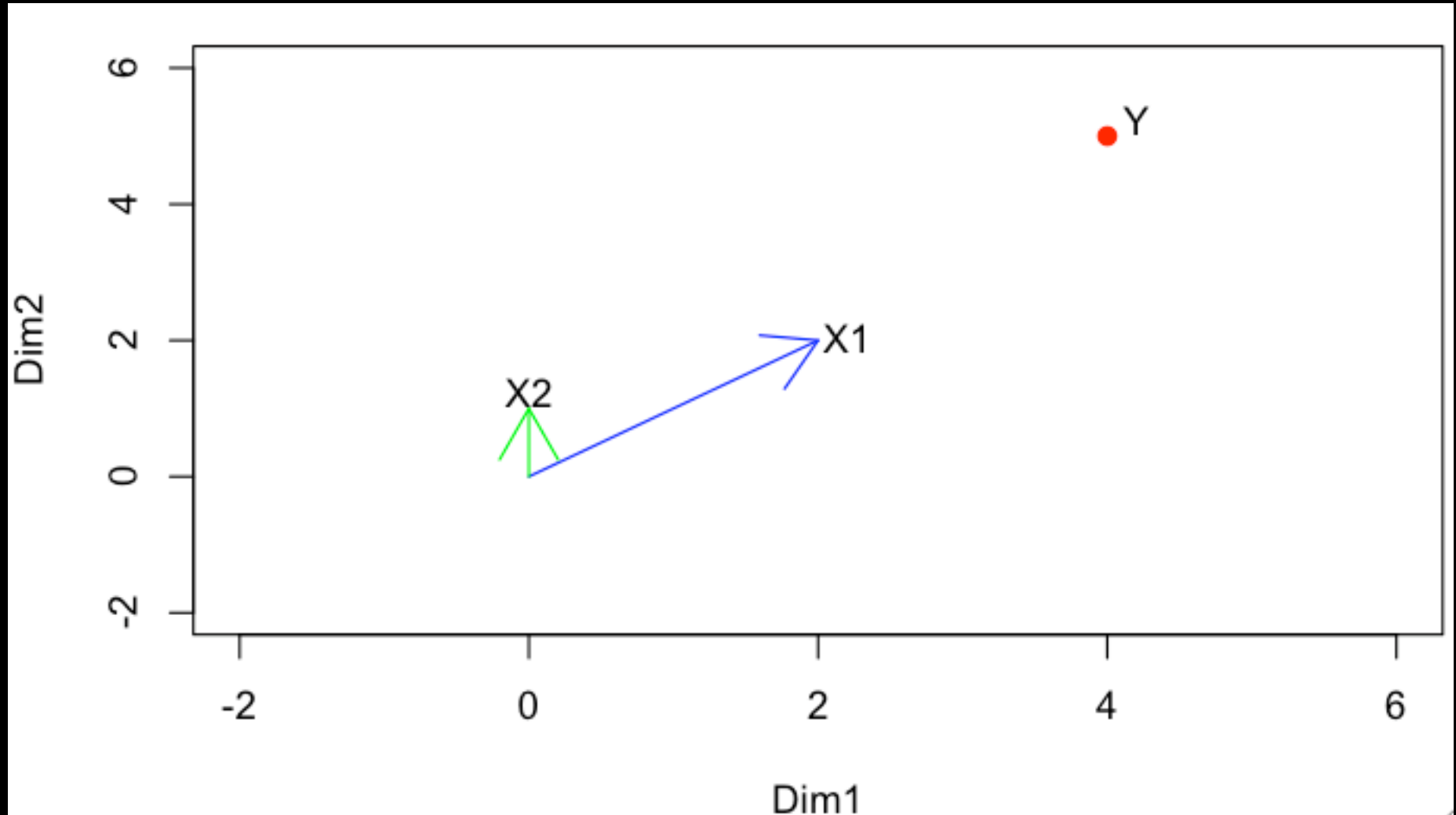
$$Y = X_1\beta_1 + X_2\beta_2$$

$$\begin{pmatrix} 4 \\ 5 \\ 0 \end{pmatrix} = \begin{pmatrix} 2 \\ 2 \\ 0 \end{pmatrix} \beta_1 + \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix} \beta_2$$

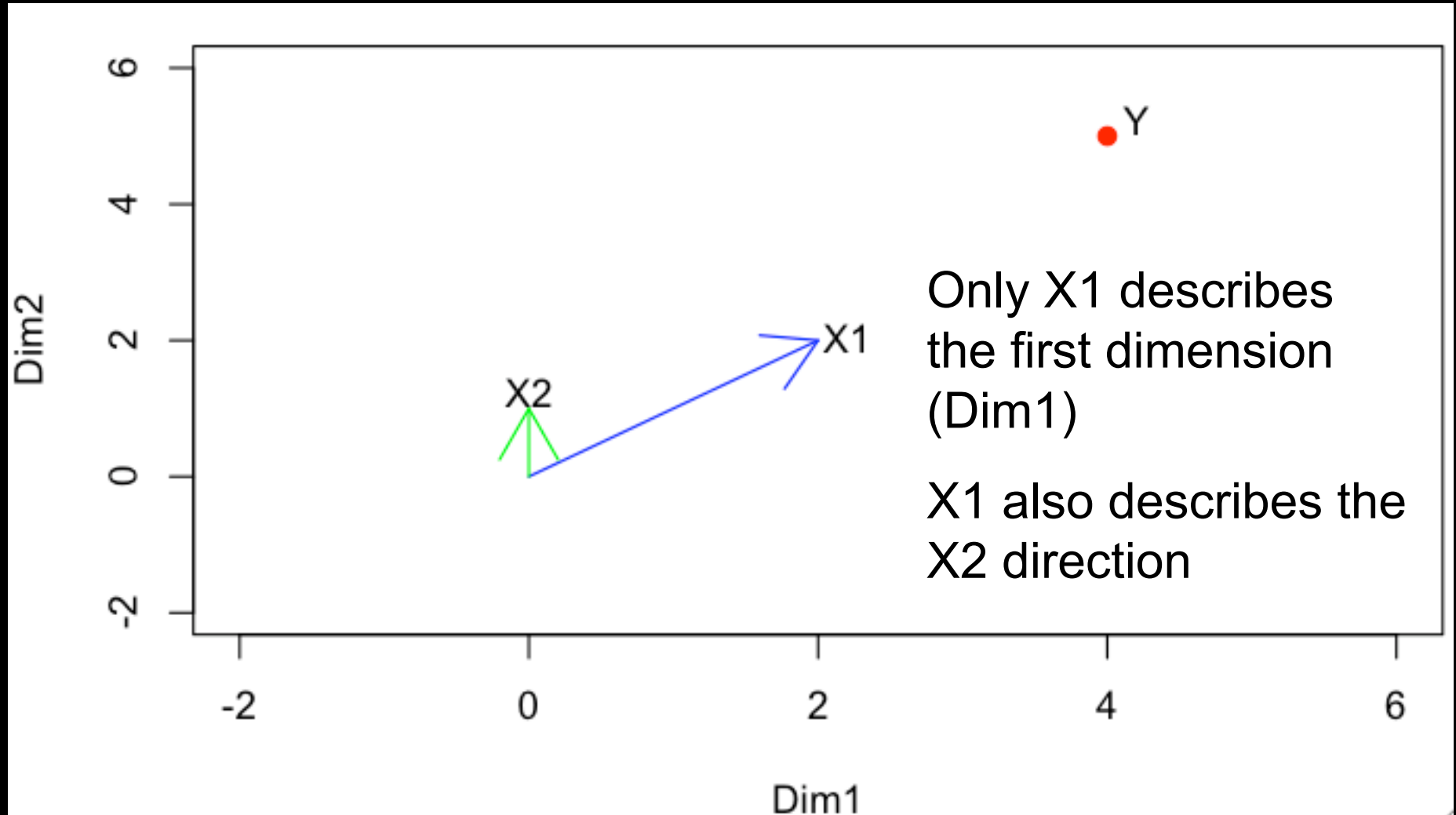
$$\beta_1=2 \text{ and } \beta_2=1$$



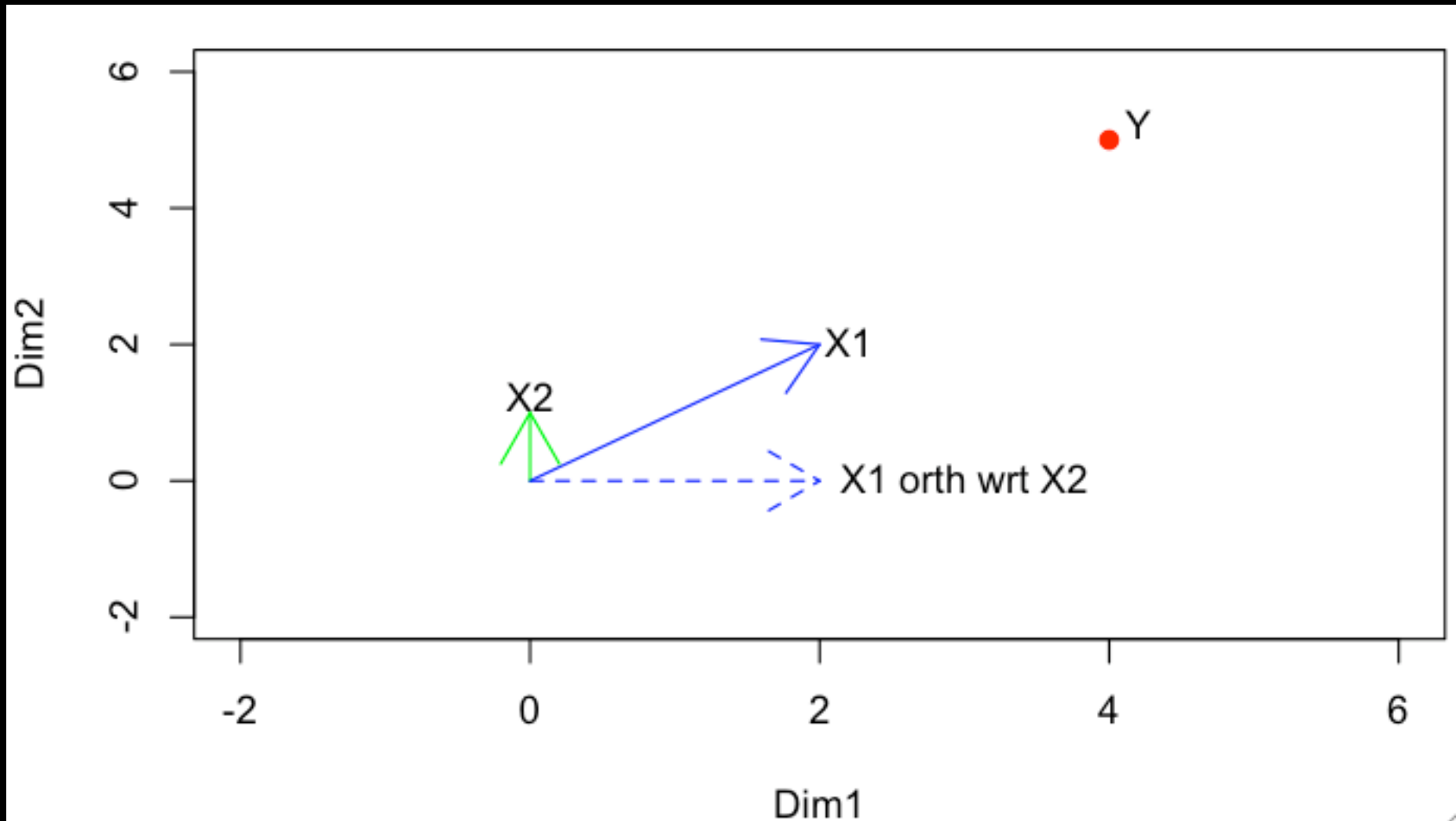
# Non-orthogonal case



# Non-orthogonal case



# Non-orthogonal case



# Non-orthogonal case

$$Y = X_1\beta_1 + X_2\beta_2$$

$$\begin{pmatrix} 4 \\ 5 \\ 0 \end{pmatrix} = \begin{pmatrix} 2 \\ 2 \\ 0 \end{pmatrix} \beta_1 + \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix} \beta_2$$

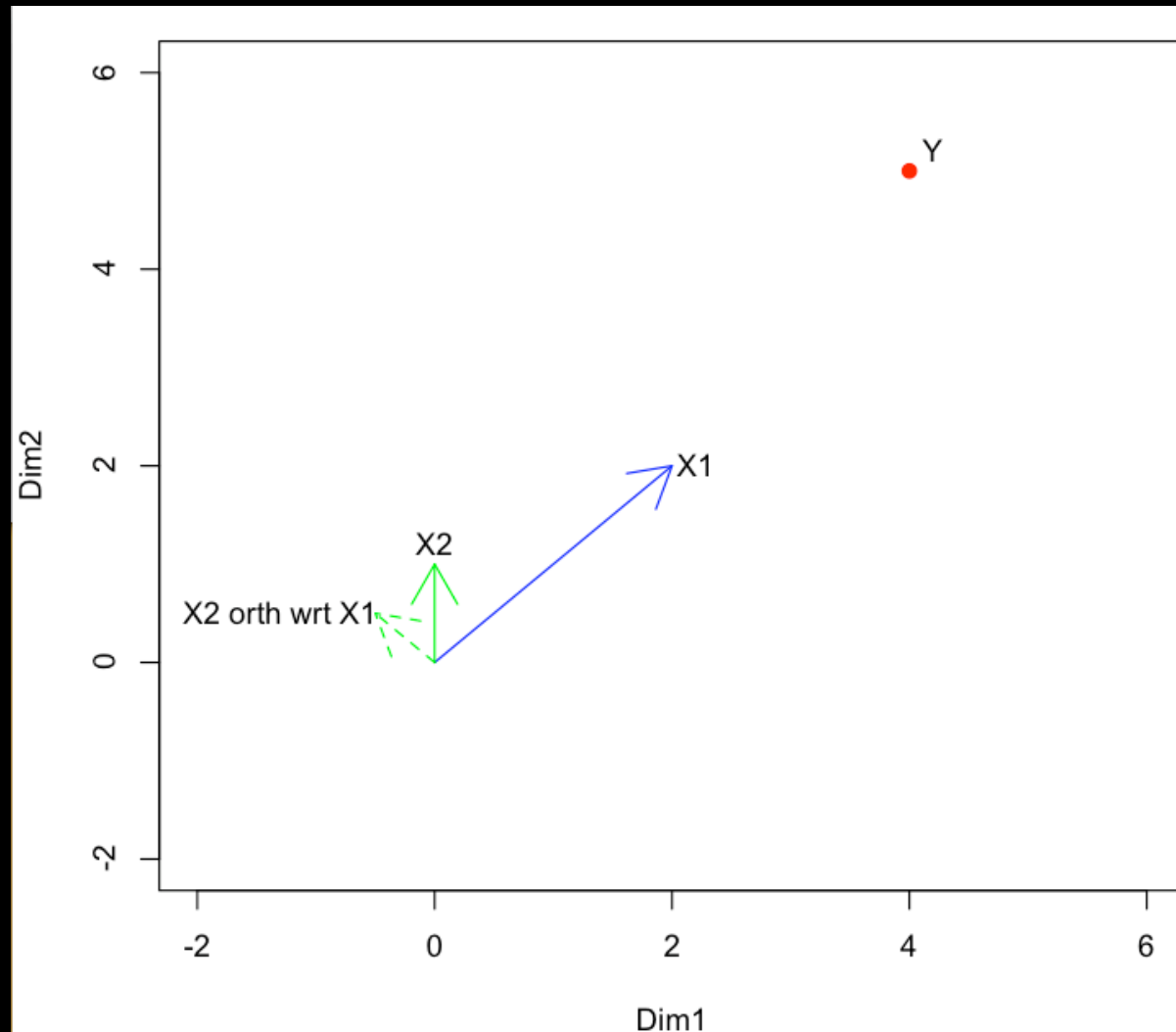
$$\beta_1=2 \text{ and } \beta_2=1$$

$$Y = X_{\text{orth}_1}\beta_1 + X_2\beta_2$$

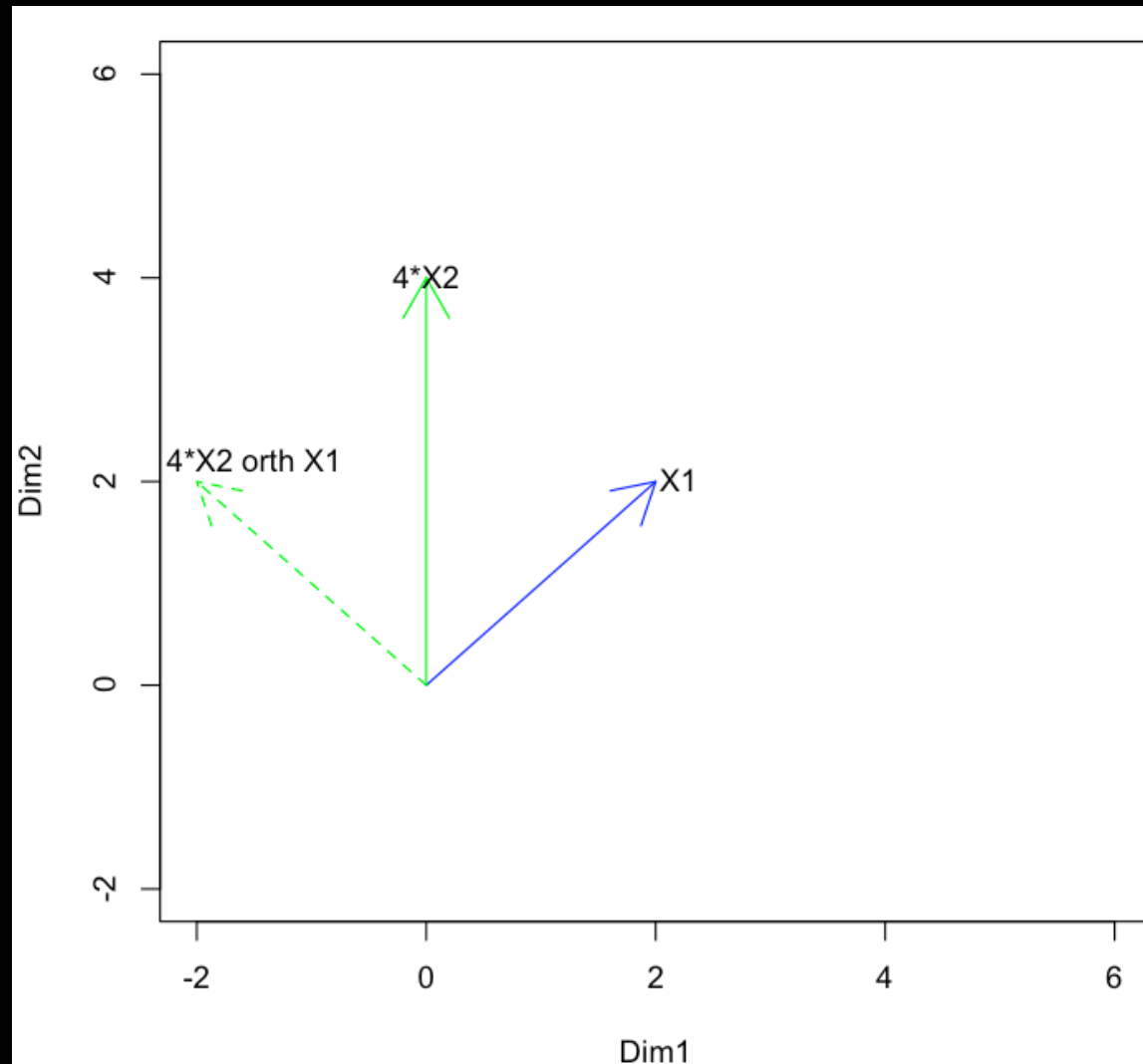
$$\begin{pmatrix} 4 \\ 5 \\ 0 \end{pmatrix} = \begin{pmatrix} 2 \\ 0 \\ 0 \end{pmatrix} \beta_1 + \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix} \beta_2$$

$$\beta_1=2 \text{ and } \beta_2=5$$

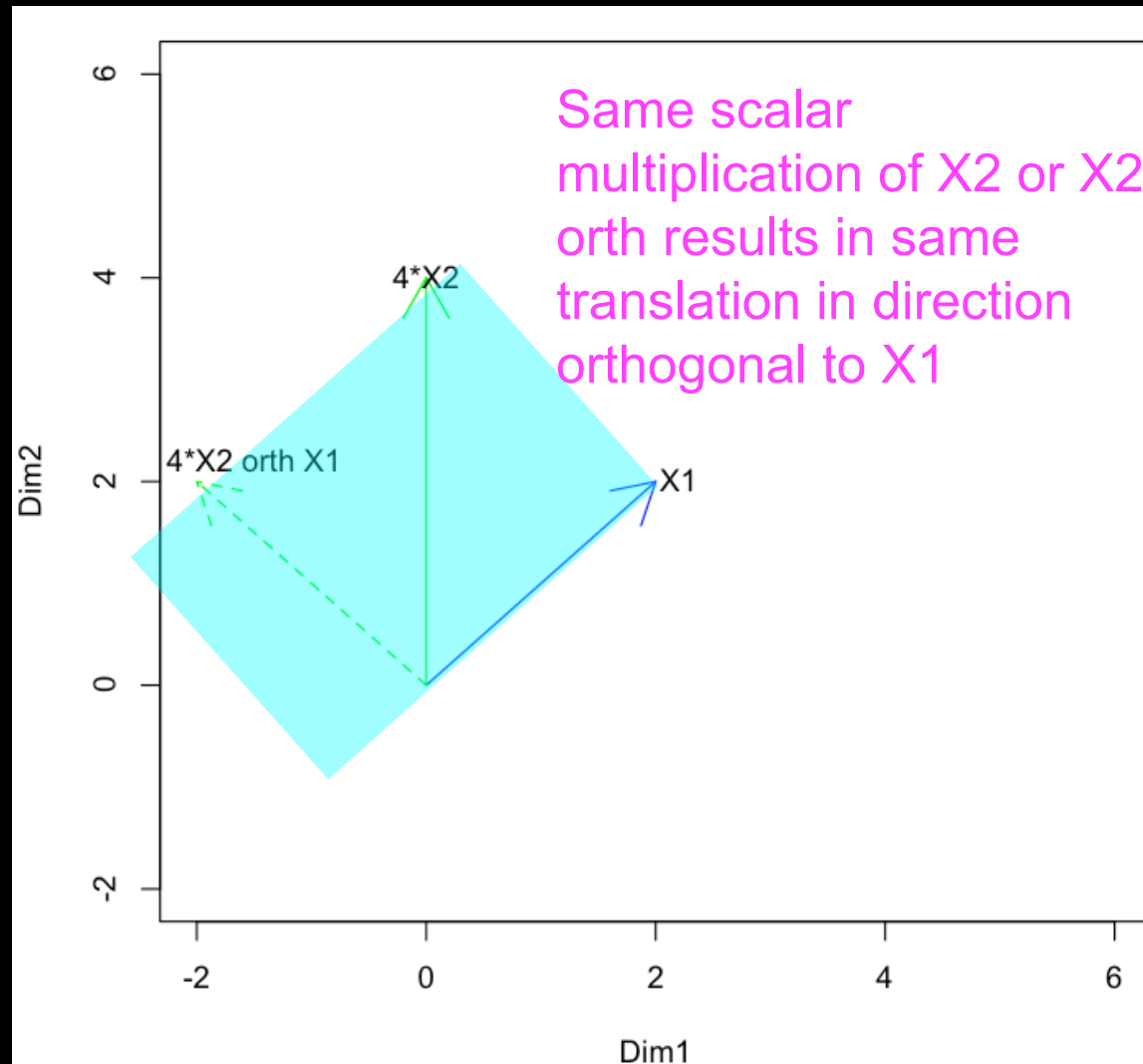
# Other way around



# Look at scalar multiplication



# Look at scalar multiplication



# Orthogonalization

- If two regressors are correlated, then they don't explain unique "directions"
- Orthogonalizing A wrt B means you are surrendering A's ability to explain variability in the B direction
- Eg, orthogonalizing age with respect to the mean is removing the ability of age to describe the overall mean of the data



# Orthogonalization

- Be careful when you interpret parameters.
  - $Y = X_1\beta_1 + X_2\beta_2$
  - Typically you interpret  $\beta_2$  as the effect of  $X_1$  adjusting for  $X_2$
  - Eg, the effect of age controlling for gender
- Orthogonalizing surrenders the ability of one covariate to control for another
  - If you model  $X_1$  orth  $X_2$ , then  $X_1$  is only soaking up extra variability **NOT** also adjusting  $X_2$

# Model 7

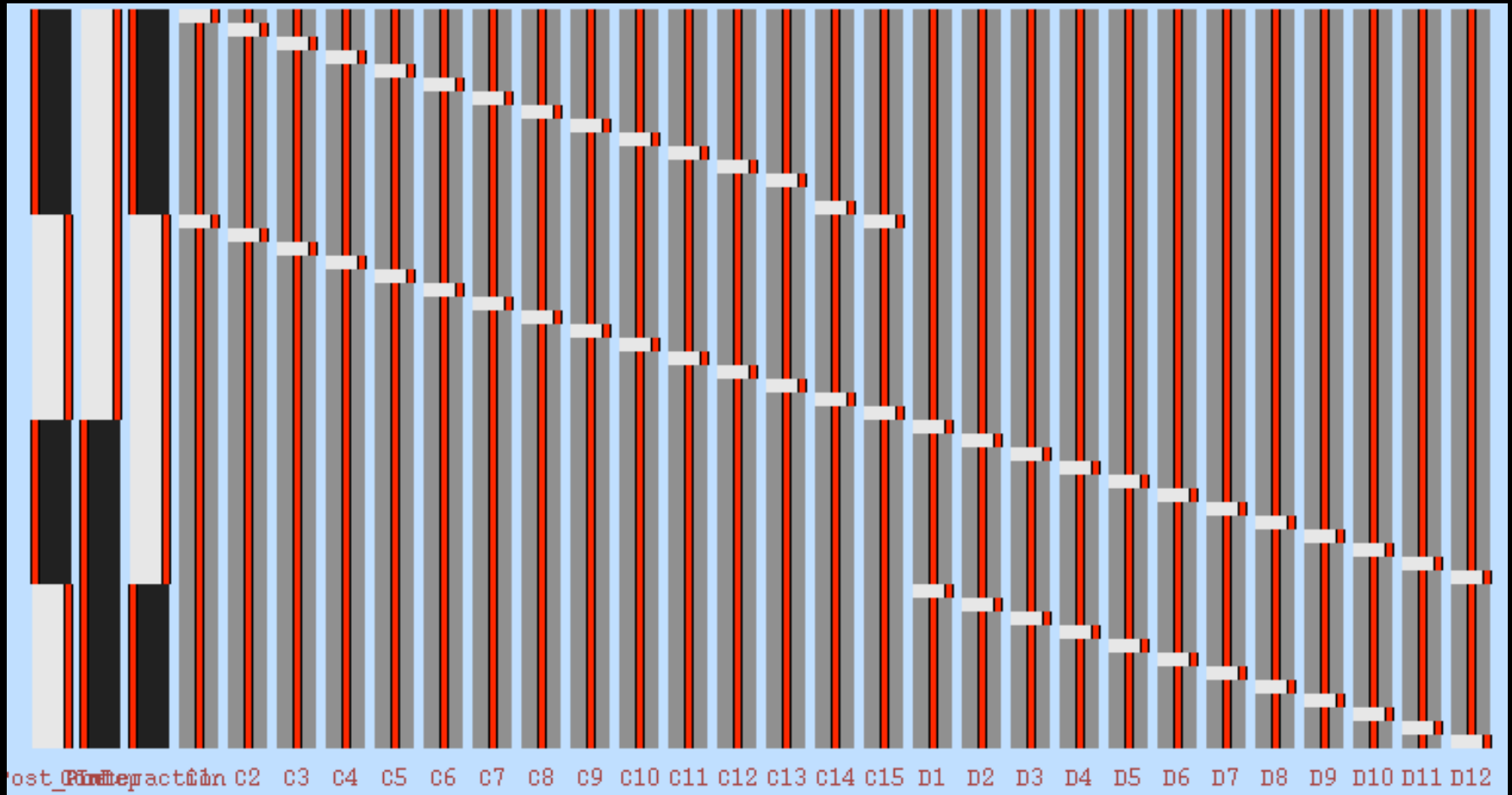
June 29, 2009, FSL help list

“We are running this model 2x2 to examine a treatment study, so all subjects (patients and controls) were scanned twice (so the model examines the GroupX Time interaction, with Time a repeated measure). We would like to examine a symptom score obtained at all scans. How would one add a symptom covariate to this model? Is it simply the addition of 4 EVs (control T1, control T2, patient T1, patient T2) of the symptom score?? If so, I assume the symptom score is demeaned separately for each of these 4 conditions (control T1, control T2, patient T1, patient T2).”

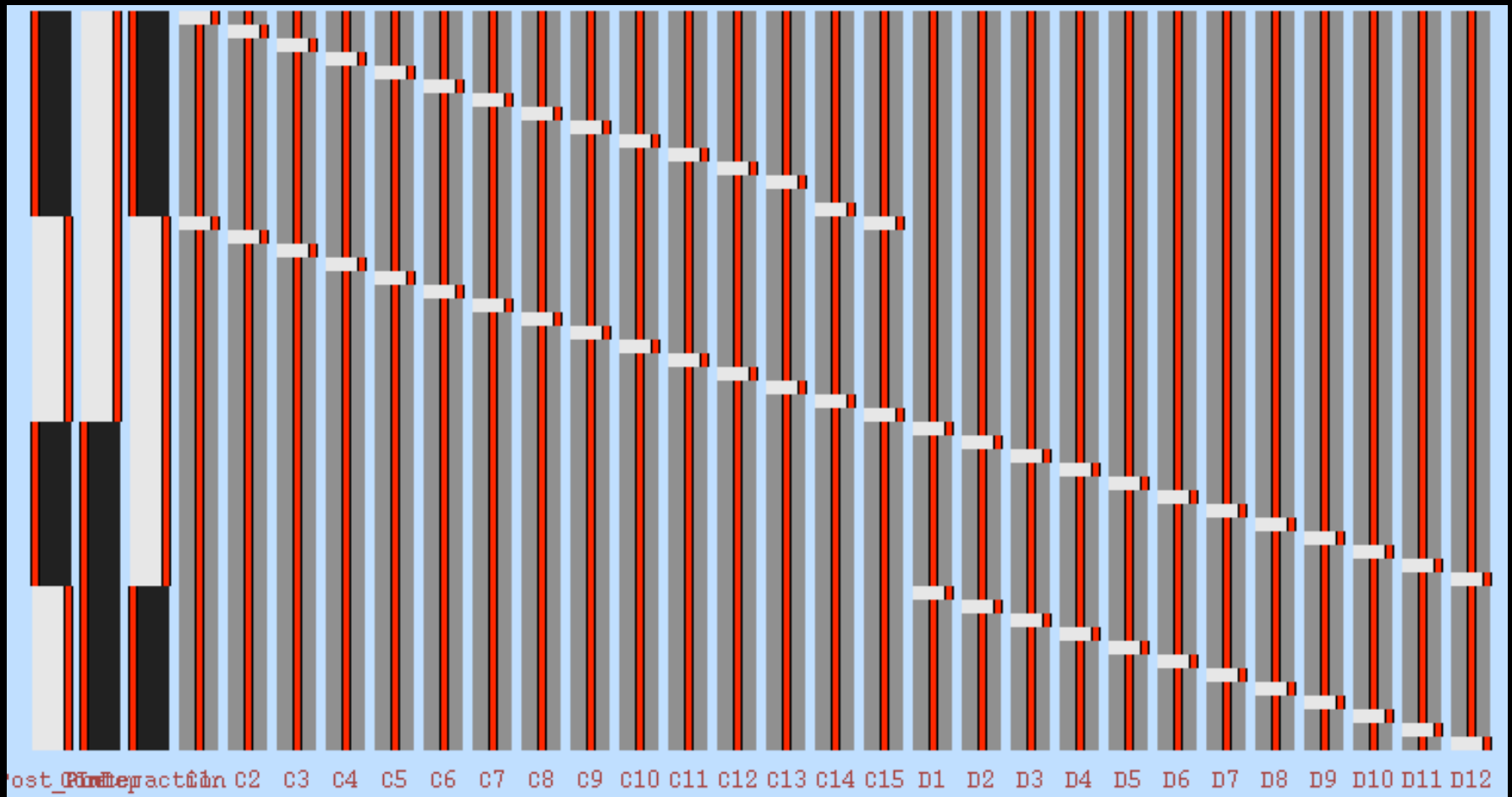
# Model 7

- Summary
  - 2x2 ANOVA with additional continuous confounder
  - Repeated measure on first factor (Time)
  - Not repeated on second factor (Group)

# What is wrong with this?



# What is wrong with this?



Rank deficient...but a good effort!



# Adding symptom covariate

- Personally, I'd first check that it didn't have structure that correlated with group and scan
- Orthogonalizing wrt ev1 and ev2 will adjust residuals for symptom, but not group or run differences
  - I'd only do this if there weren't group or run differences in symptom

Questions?